

**DARK ENERGY SPECTROSCOPIC INSTRUMENT (DESI) SURVEY YEAR 1 RESULTS**

# Validating the galaxy and quasar catalog-level blinding scheme for the DESI 2024 analysis



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**ABSTRACT:** In the era of precision cosmology, ensuring the integrity of data analysis through blinding techniques is paramount — a challenge particularly relevant for the Dark Energy Spectroscopic Instrument (DESI). DESI represents a monumental effort to map the cosmic web, with the goal to measure the redshifts of tens of millions of galaxies and quasars. Given the data volume and the impact of the findings, the potential for confirmation bias poses a significant challenge. To address this, we implement and validate a comprehensive blind analysis strategy for DESI Data Release 1 (DR1), tailored to the specific observables DESI is most sensitive to: Baryonic Acoustic Oscillations (BAO), Redshift-Space Distortion (RSD) and primordial non-Gaussianities (PNG). We carry out the blinding at the catalog level, implementing shifts in the redshifts of the observed galaxies to blind for BAO and RSD signals and weights to blind for PNG through a scale-dependent bias. We validate the blinding technique on mocks as well as on data by applying a second blinding layer to perform a series of sanity checks; the latter allows probing complexities in real data not captured in mocks. We find that the blinding strategy alters the data vector in a controlled way, and the BAO and RSD analysis choices are robust to blinding. The successful validation of the blinding strategy paves the way for the unblinded DESI DR1 analysis, alongside future blind analyses with DESI and other surveys.

**KEYWORDS:** redshift surveys, baryon acoustic oscillations, power spectrum, cosmological parameters from LSS

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**1 Introduction**

Cosmology has entered a precision era, where experiments are designed to measure key parameters of the Universe to unprecedented levels of accuracy. One of the most robust methodologies employed to understand the cosmic landscape is the two-point clustering statistics of 3D galaxy distributions. These statistics, the power spectrum  $P(k)$  in harmonic space and the correlation function  $\xi(r)$  in configuration space, provide vital clues about the underlying cosmological model and the nature of dark energy and constraints on primordial non-Gaussianities [1, 2].

However, as we refine our methods and aim for increasingly precise results, the risk of confirmation bias becomes more of a concern. These biases can arise during the data analysis process and may lead to misleading conclusions, thereby affecting the veracity of the findings. It is in this context that the concept of blind analysis becomes critically important. At the heart of it is the “blinding”, which involves the deliberate concealment or modification of key analysis outcomes, thereby ensuring that researchers’ subsequent choices and interpretations remain unbiased [3–5]. In other words, only after the full pipeline is frozen — all the choices are made —, the unaltered results are unveiled in a step we refer to as “unblinding”. Establishing clear criteria for when unblinding happens is a key part of the blind analysis procedure, which will be detailed later.

The primary goal of this paper is to present and validate the blinding technique applied to Data Release 1 (DR1; [6]) of the Dark Energy Spectroscopic Instrument (DESI) [7–9]. With its ability to collect high-quality spectroscopic data, DESI enables a range of cosmological analyses, including constraints from Baryon Acoustic Oscillations (BAO), Redshift-Space Distortions (RSD), and the investigation of scale-dependent bias due to primordial non-Gaussianities (PNG) [7, 10].

The scope of this paper is to validate the catalog-level blinding technique for galaxy and quasar samples, specifically focusing on the signals extracted from BAO and RSD analyses. While we apply the blinding scheme to PNG analyses as well, the detailed validation of the PNG-related results lies outside the current scope and will be addressed in a follow-up paper. Similarly, further optimization and validation of the BAO and RSD analyses are beyond the scope of this paper and are discussed in detail in the dedicated BAO and RSD publications [11, 12].

DESI has been designed to perform a galaxy survey spanning approximately 14,000 square degrees of the sky, encompassing regions in both the southern and northern galactic caps [13], over a period of five years. During its operation, DESI aims to determine the redshifts of around 40 million galaxies, ranging from redshifts 0.05 to 3.5. The survey has successfully completed its validation stage [14] and made its early data publicly available [15], while the analysis of DR1 is underway (for which this work is a supporting paper; more details below). DESI’s target selection program classifies its tracers into four distinct types: Bright Galaxy Survey (BGS), Luminous Red Galaxy (LRG), Emission Line Galaxy (ELG), and Quasars (QSO), in increasing order of redshift. Moreover, DESI also probes the Universe using Lyman- $\alpha$ , for which the blind analysis will follow a different type of blinding scheme; we refer the reader to [16].

We begin by giving an overview of blinding in cosmology in section 2, followed by a description of the DESI DR1 blinding scheme in section 3, expanding first on the DESI observables to motivate the parameters subjected to blinding, followed by the details of the blinding strategy; we also discuss when and under what conditions the blinding was planned to be unblinded. Next, we detail the analysis framework in section 4, discussing the data vector and covariance used, the theory model, as well as the inference framework. Then, in section 5, we validate the blinding strategy using mock datasets and the analysis framework, demonstrating the blinding technique is robust. We then validate the strategy on blinded data in section 6, delving into the statistical tests and analyses to ensure that the blinding process does not introduce any spurious features or artifacts in the data. We conclude in section 7.

## 2 Blinding in cosmology

The practice of blind analysis is not new to cosmology. In fact, different blinding strategies have been adopted for various cosmology analyses, e.g., the Supernovae analysis presented in [17] and for weak lensing surveys such as the Kilo-degree Survey (KiDS) [18–20] and the Dark Energy Survey (DES) [21]. In these applications, the blinding strategy was carefully tailored to the unique requirements and complexities of each survey and analysis. For example, the KiDS collaboration focused on the gravitational lensing signal and hence blinded their main observable, galaxy ellipticities at the catalog level [18]. On the other hand, the DES collaboration, carrying out a multi-probe experiment, employed a blinding scheme at the data-vector level ensuring internal model consistency between the galaxy clustering and weak lensing signal.

Considering the various uses of blinding for cosmology, we can distill several key criteria that a successful blinding scheme must satisfy:

- i) *Preservation of data quality*: the blinding scheme should maintain the statistical properties of the data to permit accurate validation tests.
- ii) *Difficult reversibility*: blinding should not be easily reversible by those conducting the analysis, avoiding accidental unblinding.
- iii) *Parameter specificity*: blinding should be specific to the cosmological parameters of interest, without affecting other variables and diagnostics used in the analysis. Note that the cosmological parameters of interest are defined by the actual observables the survey is most sensitive to.

These principles serve as a guideline for the choices made while developing a catalog-level blinding scheme for spectroscopic galaxy surveys in general, which we can then tailor to DESI in particular (as done in section 3.2).

The choice for catalog-level blinding is particularly motivated in order to satisfy criterion ii). The relevant quantities of a (galaxy) catalog here include two angular coordinates (right ascension, RA, and declination, DEC), the measured redshift  $z$ , and a set of weights  $w_x$  to correct for the variations in completeness. A given spectroscopic survey yields redshifts and the weights for the corresponding, pre-existing angular coordinates from the photometric catalogs from which the galaxy targets are selected. Therefore, one only needs to perturb redshifts and weights for catalog-level blinding, while leaving angular positions unchanged as these are already “unblinded” via the photometric catalog.

This procedure complies with criterion i) when it comes to the treatment of survey systematics. Due to the unchanged angular positions, the determination of systematic weights impacts the blinded and unblinded catalogs in the same fashion, hence allowing for an effective treatment of angular systematics.

Finally, to satisfy criterion iii), the exact per-object shifts in redshift and weight are not chosen to be random, and instead to distort the primary observables of a spectroscopic survey, i.e., the BAO, RSD and PNG signals. In order to ensure that the validation tests on the blinded analysis lead to insights valid on the unblinded analysis, the catalog-level blinding strategy needs to ensure that the blinded data can be represented by a viable underlying cosmological model.

### 3 DESI DR1 blinding scheme

For DESI DR1, we develop a comprehensive blinding strategy based on foundational criteria described in section 2 and methodologies described in the literature to blind for BAO and RSD [22] as well as PNG [23]. The procedure ensures that neither individual scientists nor the collective team can inadvertently unblind the data or induce experimenter biases based on intermediate results.

#### 3.1 DESI observables

The main capacity of DESI relies on examining the full 3D (along and across the line of sight) clustering of galaxies over a wide redshift range. The precise map of galaxy positions allows us to identify the cosmic web and accurately constrain the expansion history (via BAO) and growth history (via RSD) of the universe. Furthermore, the high number of large-scale modes arising from the full 3D information provides us with a powerful window to investigate the presence of PNG.

##### 3.1.1 Summary statistics

To create a 3D map of galaxies, we first convert measured galaxy redshifts  $z$  to comoving distances using a fiducial cosmological model  $\Omega_{\text{fid}}$ ; see eq. (3.13) for the fiducial model used in this paper. From the resulting catalog, we infer the galaxy redshift-space overdensity field,  $\delta_g^{\text{red}}(\vec{r})$ ; this field depends on the comoving coordinate  $\vec{r}$ . Then, we calculate the galaxy two-point clustering statistics: the correlation function  $\xi_g(s)$ , which depends on pair separation  $s$ , and its Fourier analog, the power spectrum  $P_g(k)$ , which depends on the wavevector  $k$ .

For exact implementation details of the  $\xi_g(s)$  and  $P_g(k)$  measurement, see section 4. These statistics exhibit a few distinct cosmological features described below.

##### 3.1.2 Probing the expansion history

The expansion history of the universe is encoded in the Baryonic Acoustic Oscillations (BAO) signal observed within the two-point clustering statistics. In the early universe, pressure waves, driven by the interplay between radiation and matter, propagated through the primordial plasma. This propagation continued until the universe cooled enough for protons and electrons to combine into neutral hydrogen, an epoch known as recombination. Shortly after recombination, at a redshift denoted by  $z_d$  (the drag epoch), the decoupling of baryons from photons occurred. This decoupling effectively “froze” the acoustic waves in space, marking the maximum extent to which these pressure-driven waves could travel. This maximum extent is referred to as the sound horizon and is mathematically expressed as

$$r_d = \int_{\infty}^{z_d} \frac{c_s(z)}{H(z)} dz, \quad (3.1)$$

where  $c_s(z)$  is the sound speed and  $H(z)$  the Hubble expansion rate. At late times, we can still see the impact of these waves as an overabundance of galaxy pairs and separation of the sound horizon, the BAO feature. Given that the sound horizon in comoving coordinates is fixed, it represents a *standard ruler*. Hence, by measuring the angular and parallel position of the BAO feature at different redshift bins, we can exquisitely map the angular diameter

distance  $D_A(z)$  and the Hubble distance  $c/H(z)$  in units of the sound horizon. This is often parameterized via the scaling parameters:

$$\alpha_{\perp}(z) \equiv \frac{D_A(z)r_{d,\text{fid}}}{D_{A,\text{fid}}(z)r_d}, \quad \alpha_{\parallel}(z) \equiv \frac{H_{\text{fid}}(z)r_{d,\text{fid}}}{H(z)r_d}, \quad (3.2)$$

defined with respect to a fiducial template, which is fitted to the data as detailed in section 4.2.1. The notion that the sound horizon is isotropic, i.e., its size perpendicular and parallel to the line of sight is the same, allows us to perform the so-called Alcock-Paczynski (AP) test [24]. If the fiducial cosmology chosen to transform redshifts to distances does not correspond to the underlying distance-redshift relation, this manifests itself as an anisotropy between the BAO distance perpendicular and parallel the line of sight. For this test, it is useful to combine the scaling parameters of eq. (3.2) into the isotropic (“iso”) and anisotropic (“AP”) components given as

$$\alpha_{\text{iso}} = (\alpha_{\perp}^2 \alpha_{\parallel})^{\frac{1}{3}}, \quad \alpha_{\text{AP}} = \frac{\alpha_{\parallel}}{\alpha_{\perp}}. \quad (3.3)$$

The BAO analysis hence represents a powerful tool to reconstruct the expansion history of the universe. In what follows, we consider the  $w_0w_a$ CDM model with varying dark energy equation of state

$$w(a) = w_0 + (1 - a)w_a, \quad (3.4)$$

where  $a$  is the scale factor; see [25] for a review. This parameterization allows for describing a much richer range of dynamical behavior of dark energy than a constant equation of state (i.e., the cosmological constant  $\Lambda$ ), allowing for a test of the  $\Lambda$ CDM model.

Within the flat  $w_0w_a$ CDM model, the cosmological expansion law at late times is given by

$$H(z) = H_0 \sqrt{\Omega_m(1+z)^3 + (1 - \Omega_m)(1+z)^{3(1+w_0 + \frac{z}{1+z}w_a)}}, \quad (3.5)$$

with the present time Hubble constant  $H_0$  and the relative matter energy density  $\Omega_m$ .

### 3.1.3 Probing the growth history

The growth history of the universe is mapped via the Redshift-Space Distortions (RSD) signal. By measuring galaxy redshifts, we measure galaxy velocities along the line of sight, each consisting of two contributions: the Hubble flow velocity (recession) and their own velocity (peculiar). Therefore, by converting redshifts to distances, the true real-space positions are contaminated by the peculiar velocities, giving rise to RSD. On large scales, the galaxy bulk flows can be characterized by the so-called displacement field,  $\Psi = \nabla\phi$ , the gradient of the gravitational potential  $\phi$  sourced by the real-space matter density field  $\delta_m$ . This leads to

$$\nabla \cdot \Psi = -\frac{\delta_g}{b_1}, \quad (3.6)$$

where  $b_1 = \delta_g/\delta_m$  is the scale independent linear galaxy bias. The mapping between real-space coordinate  $\vec{x}$  and redshift-space coordinate  $\vec{r}$  on large scales induced by the line-of-sight component of the displacement field  $\Psi \cdot \hat{r}$  is given, following [26], by

$$\vec{r} = \vec{x} + f(\Psi \cdot \hat{r})\hat{r} \quad (3.7)$$

with the so-called growth rate  $f \equiv d \ln D(a) / d \ln a$ , where  $D(a)$  is the linear growth function, defined as  $D(a) = \delta(a) / \delta(a = 1)$  [27]. As a consequence, the redshift-space galaxy power spectrum  $P_g(k, \mu)$  experiences an anisotropy as a function of the angle  $\mu = \cos \theta$  between the galaxy pair and the line of sight given by the Kaiser formula, [28],

$$P_g(k, \mu) \approx \left( b_1 + f \mu^2 \right)^2 P_{\text{lin}}(k), \quad (3.8)$$

where  $P_{\text{lin}}(k)$  is the linear matter power spectrum. Note that this approximation is only valid on large scales and we adopt a more sophisticated RSD model in section 4.

### 3.1.4 Probing the primordial non-Gaussianity

As for the local primordial non-Gaussianity (PNG), the primordial gravitational potential  $\phi(x)$  is approximated, as in [29], by a quadratic contribution,

$$\phi_G(x) + f_{\text{NL}} \left( \phi_G(x)^2 - \langle \phi_G(x)^2 \rangle \right), \quad (3.9)$$

where  $\phi_G$  is a Gaussian distributed random field and  $f_{\text{NL}}$  is the amplitude of the quadratic correction to the potential. The parameter  $f_{\text{NL}}$  in effect parameterizes the PNG. In the standard single-field slow-roll inflation model, the value of  $f_{\text{NL}}$  is approximately zero (it is roughly  $(n_s - 1) \simeq 0.03$ ). However, if  $f_{\text{NL}}$  is found to be substantially larger than zero, it would suggest interesting possibilities, such as the presence of multiple interacting scalar fields during the inflationary period.

## 3.2 Blinding strategy

To blind for the observables described before, we start with shifting galaxy redshifts along the line of sight, mimicking a universe with a different underlying cosmological model without changing the galaxies' angular positions. This entails two kinds of shifts: one to blind the cosmological background evolution (i.e., BAO) and another to blind the growth of structures (i.e., RSD); the first mimics the Alcock-Paczynski (AP) effect while the second mimics RSD. Both use two cosmologies: one fiducial one and one we pick for our blinding scheme.<sup>1</sup> We refer the reader to [22] for details, but summarize the two shifts briefly for completeness:

1. AP-like shift: this shift emulates the Alcock-Paczynski effect by altering galaxy positions along the line of sight. This procedure is executed by first transforming the measured redshifts  $z_i$  into comoving distances  $D_{\text{M}}(z_i, \Omega_{\text{blind}})$  using the cosmology chosen for blinding, which are then transformed back to redshift space  $z_i$  using the fiducial cosmology, as captured in equation in 3.10 [22], reproduced here for completeness:

$$z_i(\Omega_{\text{true}}) \xrightarrow{\Omega_{\text{blind}}} D_{\text{M}}(z_i, \Omega_{\text{blind}}) = D_{\text{M}}(z'_i, \Omega_{\text{fid}}) \xrightarrow{\Omega_{\text{fid}}} z'_i(\Omega_{\text{blind}}), \quad (3.10)$$

where  $\Omega_{\text{true}}$  is the cosmology underlying the observed data. As a result, once the blinded galaxy redshift catalog is converted to distances via  $\Omega_{\text{fid}}$ , the scaling parameters

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<sup>1</sup>Note that the fiducial cosmology is referred to as the “reference” cosmology in [22] while the what we call blinding cosmology is termed as “shifted” cosmology. Also, note that the arbitrary choice of fiducial cosmology does not have a sizable impact on the blinding shifts, which primarily depend on the relative difference between the fiducial and the blind cosmology.



in eq. (3.2) are shifted with respect to  $\Omega_{\text{true}}$  as predicted by  $\Omega_{\text{blind}}$ . Note that the AP-like shift is the same for all galaxies that have the same redshift.

2. RSD shift: this perturbative shift mimics the redshift-space distortions by adjusting galaxy positions based on the local galaxy density and the peculiar velocity field. The shifts in redshifts are calculated by first transforming the observed redshifts into distances using the fiducial cosmology. From the resulting galaxy redshift-space positions  $\mathbf{r}$ , and a fiducial value of the growth rate  $f^{\text{fid}}$  as input, an approximation of the real-space density and its gradient, the displacement field, are derived. Then, the line-of-sight component of the latter is used to transform each galaxy to a new, blinded redshift-space position  $\mathbf{r}'$  according to the blinding value of the growth rate  $f^{\text{blind}}$ . This is encapsulated in equation 3.18 in [22]:

$$\mathbf{r}' = \mathbf{r} - f^{\text{fid}}(\Psi \cdot \hat{\mathbf{r}})\hat{\mathbf{r}} + f^{\text{blind}}(\Psi \cdot \hat{\mathbf{r}})\hat{\mathbf{r}}. \quad (3.11)$$

As a result, the galaxy power spectrum measured from the catalog after the blinding transformation of eq. (3.11) exhibits an RSD anisotropy reminiscent to eq. (3.8) with  $f = f^{\text{true}}$  being altered to  $f = f^{\text{true}} - f^{\text{fid}} + f^{\text{blind}}$ . Note that given the impact of local galaxy density and peculiar velocity field, the shift is different for each galaxy.

The procedure described above blinds only two of the three observables that we are interested in, i.e. BAO and RSD. For the third observable, PNG, we follow the strategy in [23], which entails blinding the large scales of the power spectrum by adding weights to the data, using a blinding  $f_{\text{NL}}$  value. Here we aim to mimic the scale-dependent bias signature of PNG on large scales, which impacts the theoretical real-space galaxy power spectrum as:

$$P(k) = \left( b_1 + \frac{b_\phi}{\alpha(k)} f_{\text{NL}} \right)^2 \times P_{\text{lin}}(k), \quad (3.12)$$

where  $P_{\text{lin}}(k)$  is the linear matter power spectrum,  $\alpha(k)$  is a transfer function connecting the primordial gravitational field to the matter density perturbation,  $b_1$  is the linear bias and  $b_\phi$  is the bias quantifying the response of the tracer to local PNG.

To implement the scale-dependent part  $b_\phi f_{\text{NL}}/\alpha(k)$  of eq. (3.12) at the catalog level, we first approximately move galaxies to real-space by subtracting the RSD displacements estimated in the same manner as for the RSD shift mentioned above from the observed galaxy positions. The obtained shifted galaxies are ‘painted on a grid’<sup>2</sup> to estimate the underlying matter density field in real space  $\hat{\delta}_r$  (assuming a fiducial linear bias). The scale-dependent bias contribution  $b_\phi f_{\text{NL}} \hat{\delta}_r(\vec{k})/\alpha(k)$  is computed in Fourier space for a blinded value of  $b_\phi f_{\text{NL}}$ , then transformed back to configuration space and read off at each galaxy position, thereby providing a weight to be applied to each galaxy to mimic the scale-dependent bias.

For our blinding scheme, we pick  $w_0$ ,  $w_a$ , and  $f_{\text{NL}}$  values, as described in section 3.3. As for the fiducial cosmology, throughout this work we use Planck-2018 results [30]. The

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<sup>2</sup>*Painting galaxies on a grid* is a term commonly used to describe the process of mapping discrete galaxy positions onto a continuous grid, so that their spatial distribution can be converted into a density field.



cosmological parameters are, explicitly, given by:

$$\begin{aligned} \omega_b &= 0.02237, & \omega_{cdm} &= 0.12, & h &= 0.6736, \\ A_s &= 2.083 \times 10^{-9}, & n_s &= 0.9649, & N_{ur} &= 2.0328, \\ N_{ncdm} &= 1.0, & \omega_{ncdm} &= 0.0006442, & w_0 &= -1, \quad w_a = 0. \end{aligned} \quad (3.13)$$

$\omega_b$  and  $\omega_{cdm}$  denote the densities for baryons and cold dark matter, respectively, both scaled by  $h^2$ , where  $h$  denotes the reduced Hubble constant.  $A_s$  and  $n_s$  characterize the amplitude and spectral index of primordial scalar perturbations. Further,  $N_{ur}$  and  $N_{ncdm}$  denote the effective number of ultra-relativistic and non-cold dark matter species, with  $\omega_{ncdm}$  indicating the density of the latter. Finally,  $w_0$  and  $w_a$  are the dark energy equation of state and its evolution, as explained in section 3.1.3.

### 3.3 Blinding pipeline

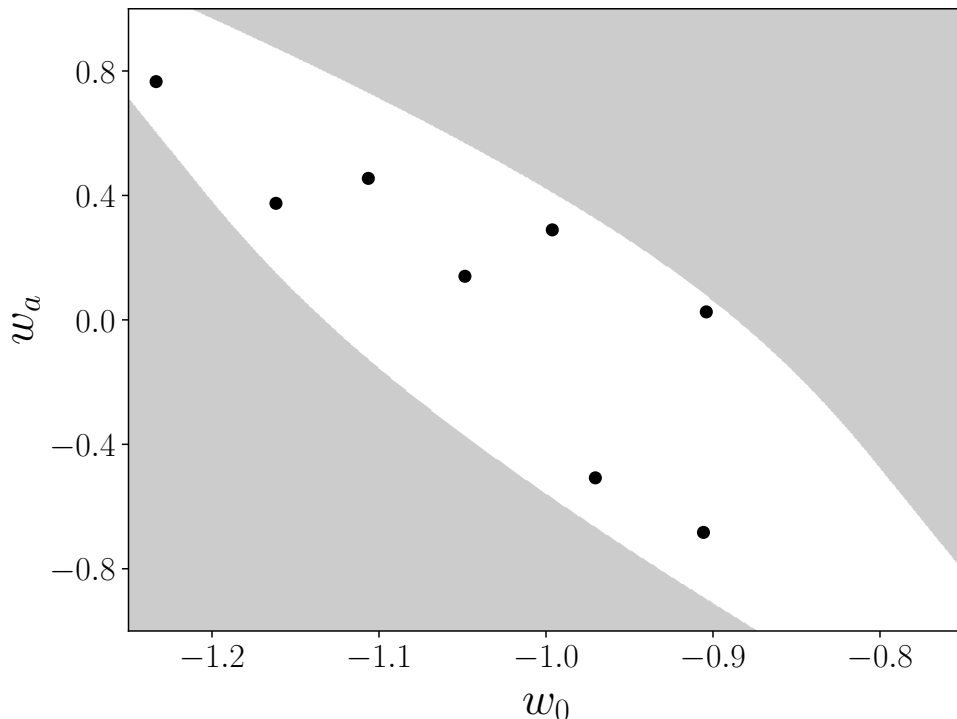
Taking into account the DESI observables discussed in section 3.1, we focus our blinding efforts on three key parameters,  $w_0$ ,  $w_a$  and  $f_{NL}$ , as they are central to the primary science goals of the DESI experiment and, thus, are highly susceptible to experimenter bias during the data validation/interpretation stage. As discussed in section 3.1, all these parameters are constrained using two-point clustering statistics from large-scale structure observations, with details of the analysis framework in section 4. While it would be possible to further extend the blinding parameter basis, for example, by adding non-zero curvature, we decided to limit the AP blinding to the flat  $w_0 w_a$ CDM model introduced in section 3.1, since with DESI alone we do not expect to constrain the dark energy equation of state jointly with curvature, due to the strong degeneracy of these parameters. This means, that given the DESI precision, blinding for flat  $w_0 w_a$ CDM model imprints sufficient freedom to the  $H(z)$  function that makes it barely indistinguishable from a  $H(z)$  function of a  $k$ -CDM model.

To ensure the robustness of our blinding scheme, we confine the shifts in the blinded cosmology to specific regions within the  $(w_0, w_a)$  parameter space. This allows us to ensure that these shifts can be accurately translated into galaxy redshift changes. In particular, we dictate that the shifts for the BAO-scaling parameters  $\alpha_\perp$  and  $\alpha_\parallel$ , defined in eq. (3.2), should be kept within a maximum deviation of 3% from their fiducial value of unity, i.e.,  $|\alpha_\perp - 1| < 0.03$  and  $|\alpha_\parallel - 1| < 0.03$ , respectively. Figure 1 illustrates these constraints, showcasing a  $(w_0, w_a)$  region permissible within the redshift range  $0.4 < z < 2.1$  (white region), from which we pick  $(w_0, w_a)$  values for blinding; details of the redshift range used for this selection are given in appendix A. Also, we require that the amplitude of the monopole of the clustering signal does not change significantly, i.e., we aim to keep it as close as reasonably possible to the true one.<sup>3</sup> We compute<sup>4</sup> the impact of the blinded cosmology on the growth factor  $f$  as

$$f_{\text{blind}}(z) = b_1(z) \left( \sqrt{\frac{D_{A,\text{fid}}^2(z) H_{\text{blind}}(z, w_0, w_a)}{D_{A,\text{blind}}^2(z, w_0, w_a) H_{\text{fid}}(z)}} \left( \frac{f_{\text{fid}}^2(z)}{b_1^2(z)} + \frac{10}{3} \frac{f_{\text{fid}}(z)}{b_1(z)} \right) + \frac{25}{9} - \frac{5}{3} \right), \quad (3.14)$$

<sup>3</sup>For a definition of multiples moments see section 4.1.

<sup>4</sup>eq. (3.14) is derived by requiring the change in power spectrum monopole amplitude due to RSD from eq. (3.8) to compensate the volume dilation factor proportional to  $\alpha_{\text{iso}}^{-3}$  arising from the AP blinding.



**Figure 1.** Parameter space of interest for  $(w_0, w_a)$  under the DESI DR1 blinding scheme. The white region represents the parameter region that allows for changes in  $\alpha_{\parallel}$  and  $\alpha_{\perp}$  of less than 3% with respect to a fiducial chosen value of 1 in the redshift range  $0.4 < z < 2.1$ . The black points are 8 random selections used to blind our mock catalogs, which we use to validate our methodology.

and require that the shifts in  $f$  do not exceed 10% of the fiducial value,  $f_{\text{fid}} = 0.8$ . These ranges (3% for  $\alpha_{\perp}$  and  $\alpha_{\parallel}$ , 10% for  $f$ ) were roughly based on the precision of such measurements before DESI DR1.

For validating the blinding scheme, we randomly select 8 pairs of  $(w_0, w_a)$ , shown as black dots in figure 1, as well as two  $f_{\text{NL}}$  values ( $\pm 20$ ), to blind our mock catalogs. This validation is described in section 5.

To blind the data catalogs, we generate a list of 1,000 random combinations of  $(w_0, w_a)$ , all within the white region of figure 1. Then, from the list of 1,000 pairs of  $(w_0, w_a)$ , we randomly select one to blind our data catalog (using the same pair of values for all our tracers), following the prescription in section 3.2. We do not disclose the parameters used for blinding the data given that not all the DESI DR1 papers are unblinded as of this writing; note that the same blinding is used for BAO, RSD, and  $f_{\text{NL}}$  DESI analyses. Importantly, even the ‘blinding team’ did not have access to the specific blinding parameters. This ensures that ongoing projects, such as the  $f_{\text{NL}}$  and RSD validations, remain securely blinded despite the unblinding of BAO results.

By adhering to these principles, our DESI DR1 blinding scheme aims to provide a reliable and effective means to mitigate the confirmation bias effect, thereby ensuring the integrity of the DESI DR1 cosmology analysis.

### 3.4 When to unblind: criteria and tests

The decisions on the blinding pipeline were coordinated closely with the DR1 DESI collaboration 2-point clustering cosmological analysis team. Once milestones for analysis validation on blinded catalogs were reached;<sup>5</sup> the results were unblinded. The results particularly relevant to this work are presented in:

1. DESI 2024 II: sample definitions, characteristics, and two-point clustering statistics [31]
2. DESI 2024 III: Baryon Acoustic Oscillations from Galaxies and Quasars [11]
3. DESI 2024 V: analysis of the full shape of two-point clustering statistics from galaxies and quasars [12]

We emphasize once more that the scope of the present paper is to ensure that the blinding pipeline works as expected (as detailed in the next section), and we refer the reader to [11, 12] for details regarding the respective analyses, including optimizations and unblinding requirement tests, which are beyond the scope of the present paper. We also note that DESI 2024 IV focuses on BAO measurements Lyman- $\alpha$  forest; blinding, validation, and analysis for IV were distinct from those described above; we refer the reader to [16]. Finally, DESI 2024 VI focuses on likelihoods and combining the various probes; see [32, 33] for further details.

### 3.5 Blinding pipeline validation

As previously mentioned, the primary objective of this paper is to demonstrate that the blinding pipeline performs as intended. Specifically, we validate the pipeline by applying it to both mock and real (blinded) data (i.e., section 5, section 6), *ensuring that we observe only the expected shifts deliberately introduced by the blinding procedure, and no other remaining artifact*. These checks confirm that the blinding process is effective, without unintentionally revealing any cosmological information prior to unblinding, and that no uncontrolled deviations occur. Given that we understand how the blinding affects our measurements, we can safely utilize the catalogs that underwent through the pipeline and then study systematic error budgets. After those errors are set, one can finally unblind.

## 4 Analysis framework

Here we elaborate on how we calculate the two-point statistics from the catalog (our data vectors). We also explain the theoretical model as well as the compression employed to extract information from the two-point statistics, followed by the details of the inference framework that we adopted.

### 4.1 Data vector and covariance

#### 4.1.1 Correlation function estimator

In order to measure the two-point correlation function, which measures the excess probability of finding two galaxies at a specific separation  $s$  and angle relative to the line-of-sight  $\mu$ ,

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<sup>5</sup>For a detailed list of milestones reached in the BAO analysis, see section 6 of [11]. A similar list of unblinding criteria and tests for the full-shape analysis will be available in the forthcoming full-shape paper [12].

we use the Landy-Szalay estimator [34],

$$\hat{\xi}(s, \mu) = \frac{DD(s, \mu) - DR(s, \mu) - RD(s, \mu) + RR(s, \mu)}{RR(s, \mu)}, \quad (4.1)$$

where  $DD$ ,  $DR$ ,  $RD$ , and  $RR$  represent the weighted number of data ( $D$ ) and random ( $R$ ) pairs in specific distance and angle bins. From this, we calculate the multipole moments (monopole given by  $\ell = 0$ , quadrupole by  $\ell = 2$ , and hexadecapole by  $\ell = 4$ ) using Legendre polynomials,

$$\hat{\xi}_\ell(s) = \frac{2\ell + 1}{2} \int d\mu \hat{\xi}(s, \mu) \mathcal{L}_\ell(\mu). \quad (4.2)$$

#### 4.1.2 Power spectrum estimator

The power spectrum estimator, on the other hand, leverages the Feldman-Kaiser-Peacock (FKP) estimator [35, 36], which combines galaxy and random field densities to calculate the weighted galaxy fluctuation field,

$$F(\vec{r}) = n_d(\vec{r}) - \alpha n_r(\vec{r}), \quad (4.3)$$

where  $n_d(\vec{r})$  and  $n_r(\vec{r})$  are the weighted galaxy and random number densities, the latter having a total weighted number  $1/\alpha$  times the one of the data catalog. For a detailed discussion of the weighting scheme, we refer to [31].

Power spectrum multipoles are then calculated as an average over all Fourier modes  $\vec{k}$  within a bin centered on the magnitude  $k$ . The sum over  $\vec{k}$  corresponds to summing over the different wavevectors that fall within the bin. The estimator is given by:

$$\hat{P}_\ell(k) = \frac{2\ell + 1}{AN_k} \sum_{\vec{k} \in k} \sum_{\vec{r}_1} \sum_{\vec{r}_2} F(\vec{r}_1) F(\vec{r}_2) \mathcal{L}_\ell(\hat{k} \cdot \hat{\eta}) e^{i\vec{k} \cdot (\vec{r}_2 - \vec{r}_1)} - \mathcal{N}_\ell, \quad (4.4)$$

where the sums are performed over galaxy pairs with positions  $\vec{r}_1, \vec{r}_2$  and line-of-sight  $\hat{\eta}$ , and over wavevectors  $\vec{k}$  within the bin of magnitude  $k$ .  $\mathcal{N}_\ell$  denotes the shot noise correction applied to the monopole term, and  $A$  is the normalization factor.  $N_k$  is the number of modes in the  $k$ -bin.

#### 4.1.3 Measurements

We use `pycorr`<sup>6</sup> and `pypower`<sup>7</sup> to execute the two estimators above. As for covariances, we utilize those generated with the `RascalC`<sup>8</sup> [39] code for configuration space, and those from `TheCov`<sup>9</sup> [40, 41] code for Fourier space.

<sup>6</sup><https://github.com/cosmodesi/pycorr>: `pycorr` is essentially a wrapper of a modified version of the `CORRFUNC` package [37].

<sup>7</sup><https://github.com/cosmodesi/pypower>: `pypower` is based on the original `NBODYKIT` [38] implementation.

<sup>8</sup><https://github.com/oliverphilcox/RascalC>.

<sup>9</sup><https://github.com/cosmodesi/thecov>.

## 4.2 Compression approaches and theory models

We compress the information encoded in the two-point statistics, allowing fitting the two-point function to a *template* with only a limited set of parameters. Before we delve into the fitting method, we summarize three approaches to compress information from the two-point clustering statistics, in increasing complexity:

1. *Standard BAO approach*: this method focuses on extracting the cosmological distance scale from the BAO observed in both pre- and post-reconstructed correlation functions and power spectra; see e.g. for reference [42]. Specifically, it utilizes the isotropic ( $\alpha_{\text{iso}}$ ) and anisotropic ( $\alpha_{\text{AP}}$ ) dilation scales to infer the Hubble parameter and angular diameter distance relative to the sound horizon at the drag epoch. While extracting the BAO feature from clustering statistics, a polynomial expansion is often used to parameterize the broadband, allowing us to marginalize over non-BAO peak information. This approach provides a robust means to measure the expansion history of the Universe [43–45].
2. *Standard BAO+RSD approach*: this approach extends the standard BAO analysis by incorporating measurements of RSD, enabling the extraction of the growth rate of structure parameter ( $f$ ) alongside the geometric BAO signals in the pre-reconstructed catalogs. The combined analysis not only enhances the constraining power on cosmological parameters, particularly those related to dark energy and gravity theories, but also measures the rate of gravitational clustering through the  $df$  parameter, providing a direct probe of the theory of gravity on cosmic scales [46]. The set of parameters constrained then is  $\{\alpha_{\text{iso}}, \alpha_{\text{AP}}, df\}$ , where  $df$  captures the rate of structure formation. (In our framework,  $df$  is defined as the ratio  $f/f^{\text{fid}}$ , and is calculated using the `desilike` package for model predictions.<sup>10</sup>)
3. *ShapeFit*: applied exclusively to pre-reconstructed power spectra, ShapeFit goes beyond BAO and RSD by incorporating an additional parameter  $m$  to model the broadband shape of the power spectrum. The expanded set of parameters:  $\{\alpha_{\text{iso}}, \alpha_{\text{AP}}, m, df\}$  allows for a more comprehensive analysis of the cosmic expansion history and growth of structure, leveraging the complementary information encoded in the shape of the power spectrum and the amplitude of RSD [47, 48].

In this paper, we utilize the standard BAO approach and ShapeFit to capture the two ends of complexity for our validation tests. Now we turn to explaining how exactly the two compression analyses work.

### 4.2.1 Standard BAO compression

The standard approach utilizes a pre-defined template based on theoretical predictions that capture the effects of BAO and a broadband term that marginalizes over RSD in the clustering of galaxies. By adjusting the template’s amplitude, scale, and shape to best match the observed data, we can infer distances. The fitting template for the power spectrum is defined, as in [58], as

$$P(k, \mu) = \mathcal{B}(k, \mu)P_{\text{nw}}(k) + \mathcal{C}(k, \mu)P_{\text{w}}(k) + \mathcal{D}(k), \quad (4.5)$$

<sup>10</sup><https://github.com/cosmodesi/desilike>.

where  $P_{\text{nw}}(k)$  is the smooth (no-wiggle) component of the linear power spectrum and  $P_{\text{w}}(k)$  is the BAO (wiggle) component. Both components are obtained using the *peak average* method from [49].

The term  $\mathcal{B}(k, \mu)$  incorporates both the RSD and linear bias, while  $\mathcal{C}(k, \mu)$  extends this by including an anisotropic damping factor that describes the broadening of the BAO signal due to large-scale bulk flows. This damping is characterized by the parameter  $\Sigma$ , which accounts for the scale-dependent distortion of the BAO signal. The  $\Sigma$  parameters include contributions from both the real-space distortions and the effects of small-scale velocities, commonly known as the Fingers-of-God (FoG) effect. Recent work by [58] suggests that the FoG damping should be applied exclusively to the smooth component,  $P_{\text{nw}}(k)$ , which is the approach adopted here and in [11].

Finally,  $\mathcal{D}(k)$  accounts for any residual deviations from linear theory in the broadband shape of the power spectrum multipoles. The model in eq. (4.5) is then integrated over  $\mu$  to produce predictions for the power spectrum multipoles:

$$P_\ell(k) = \frac{2\ell+1}{2} \int_{-1}^1 d\mu \mathcal{L}_\ell(\mu) [\mathcal{B}(k, \mu) P_{\text{nw}}(k) + \mathcal{C}(k'(k, \mu), \mu'(k, \mu)) P_{\text{w}}(k'(k, \mu))] + \mathcal{D}_\ell(k). \quad (4.6)$$

The term involving BAO wiggles is evaluated at  $k'$  and  $\mu'$ , which are given by

$$k'(k, \mu) = \frac{k}{\alpha_\perp} \sqrt{1 + \mu^2 \left( \frac{\alpha_\perp^2}{\alpha_\parallel^2} - 1 \right)} \quad (4.7)$$

and

$$\mu'(\mu) = \frac{\mu}{\frac{\alpha_\parallel}{\alpha_\perp} \sqrt{1 + \mu^2 \left( \frac{\alpha_\perp^2}{\alpha_\parallel^2} - 1 \right)}}, \quad (4.8)$$

where  $\alpha_\perp$  and  $\alpha_\parallel$  are the BAO scaling parameters across and along the line of sight, respectively, defined in eq. (3.2). The measured  $\alpha_\perp$  and  $\alpha_\parallel$  can be transformed into the isotropic and anisotropic BAO dilations  $\alpha_{\text{iso}}$  and  $\alpha_{\text{AP}}$  provided in eq. (3.3). The latter represents the parameter basis we use throughout the rest of this work.

#### 4.2.2 ShapeFit compression

This approach incorporates additional shape information from the galaxy power spectrum, while also fitting for BAO and RSD features.

Within the ShapeFit formalism the scale dependence of the linear power spectrum  $P_{\text{lin}}$  is represented by the following modification of the fiducial template  $P_{\text{lin}}^{\text{fid}}$  via the shape parameter  $m$

$$P_{\text{lin}}(k) = P_{\text{lin}}^{\text{fid}}(k) \exp \left\{ \frac{m}{a_m} \tanh \left[ a_m \ln \left( \frac{k}{k_p} \right) \right] \right\}. \quad (4.9)$$

Here,  $k_p = \pi/r_d$  is the pivot scale and  $a_m = 0.6$  is tuned to fit the full numerical calculation of the linear power spectrum with a Boltzmann code such as CLASS [50] or CAMB [51] over a wide model parameter space; see [47] and in particular figure 4 therein for reference.

Tracer	Redshift	Analysis Type	$k_{\text{lim}}$ ( $h/\text{Mpc}$ )	$s_{\text{lim}}$ ( $\text{Mpc}/h$ )
BGS	[0.1, 0.4]	BAO	[0.02, 0.3]	[50, 150]
		ShapeFit	[0.02, 0.2]	[32, 150]
LRG	[0.4, 0.6], [0.6, 0.8], [0.8, 1.1]	BAO	[0.02, 0.3]	[50, 150]
		ShapeFit	[0.02, 0.2]	[30, 150]
ELG	[0.8, 1.1], [1.1, 1.6]	BAO	[0.02, 0.3]	[50, 150]
		ShapeFit	[0.02, 0.2]	[27, 150]
QSO	[0.8, 2.1]	BAO	[0.02, 0.3]	[50, 150]
		ShapeFit	[0.02, 0.2]	[25, 150]

**Table 1.** Summary of galaxy tracers, their redshift ranges, and the applied scale cuts for BAO analysis ( $\ell = 0, 2$ ) and ShapeFit ( $\ell = 0, 2, 4$ ).

Underpinning our ShapeFit analysis is the Lagrangian Perturbation Theory (LPT) approach to large-scale structure, provided by the `velocileptors`<sup>11</sup> code [52]. This tool computes the redshift-space distortions and clustering statistics using perturbation theory, including non-linearities crucial for accurate modeling at scales smaller than  $k \geq 0.07 h \text{ Mpc}^{-1}$ . We refer the reader to [52] for the detailed modeling.

### 4.3 Inference framework

We apply our analysis pipeline to extract key cosmological information from the two-point statistics, using theoretical models and data explained earlier, summarizing the complex data into a few interpretable parameters. We implement a series of scale cuts as detailed in table 1 for analyses both in Fourier and configuration space, specifically, the correlation function multipoles are obtained as a Hankel-transform of the power spectrum multipoles. Finally, the measured power spectrum in relation to theoretical models is mediated through the window function, incorporating survey geometry and selection effects, using the formalism detailed in [53].

We employ a Bayesian inference framework to extract the compressed parameters from galaxy correlation and power spectrum measurements, implemented in the `desilike` framework. To sample the posteriors, we utilize the Markov Chain Monte Carlo (MCMC) method implemented in the `emcee` package, with Gelman-Rubin convergence diagnostic of  $R-1 < 0.02$ . We also require the effective sample size of the chains<sup>12</sup> to be  $\gtrsim 10^3$ . Additionally, we do profile likelihood with `iminuit` package.<sup>13</sup> The priors used are included in table 2.

### 4.4 Blinding validation metrics: definition of $\Gamma$

In this section, we describe how we quantify any net residual (i.e., bias) in the measured dilation and ShapeFit parameters after accounting for the expected effect of the shift introduced

<sup>11</sup><https://github.com/sfschen/velocileptors>.

<sup>12</sup>Defined as the maximum of the weighted chain length divided by the autocorrelation length for all parameters.

<sup>13</sup><https://github.com/scikit-hep/iminuit>.



Parameter	Prior	Description
BAO Template		
$\alpha_{\text{iso}}$	$\mathcal{U}(0.8, 1.2)$	Isotropic distortion parameter
$\alpha_{\text{AP}}$	$\mathcal{U}(0.8, 1.2)$	Alcock-Paczynski distortion parameter
ShapeFit Template		
$\alpha_{\text{iso}}$	$\mathcal{U}(0.8, 1.2)$	Isotropic distortion parameter
$\alpha_{\text{AP}}$	$\mathcal{U}(0.8, 1.2)$	Alcock-Paczynski distortion parameter
$m$	$\mathcal{U}(-3, 3)$	Shape parameter
$df$	$\mathcal{U}(0, 2)$	Growth rate parameter
<b>velocileptors</b> Theory		
$b_1$	$\mathcal{U}(-1, 10)$	Linear bias, density relation.
$b_2$	$\mathcal{N}(0, 10^2)$	Second-order bias, non-linear effects.
$b_s$	$\mathcal{N}(0, 5^2)$	Tidal bias, anisotropic clustering.
$\alpha_0$	$\mathcal{N}(0, 30^2)$	Monopole shot noise.
$\alpha_{2(4)}$	$\mathcal{N}(0, 50^2)$	Quadrupole (and hexadecapole) shot noise.
$s_{n,0}$	$\mathcal{N}(0, 4^2)$	Monopole stochastic term.
$s_{n,2}$	$\mathcal{N}(0, 100^2)$	Quadrupole stochastic term.
$s_{n,4}$	$\mathcal{N}(0, 500^2)$	Hexadecapole stochastic term.

**Table 2.** Parameter priors and descriptions for the BAO and ShapeFit templates as well as **velocileptors** used in our analysis. We note that when excluding the hexadecapole from our analysis, we set  $\{\alpha_4, s_{n,4}\} = 0$ . Moreover, following the **velocileptors** paper [52], the  $b_3$  parameter, representing third-order bias, is set to zero throughout this work.

by the blind cosmologies. We first rescale the measured values with a given cosmology in terms of the baseline cosmology distance ratios as

$$\alpha^{\text{measured}} \rightarrow \frac{\alpha^{\text{measured}}}{\alpha^{\text{rescaling}}} \quad (4.10)$$

with

$$\alpha_{\text{iso}}^{\text{rescaling}} = \frac{D_V^{\text{baseline}}}{D_V^{\text{fid}}} \frac{r_d^{\text{fid}}}{r_d^{\text{baseline}}}, \quad \alpha_{\text{AP}}^{\text{rescaling}} = \frac{D_M^{\text{baseline}}/D_H^{\text{baseline}}}{D_M^{\text{fid}}/D_H^{\text{fid}}}, \quad (4.11)$$

and where  $D_V \equiv (D_H(z)D_M(z)^2)^{1/3}$  is the spherically averaged distance. The baseline here will be **a** blind cosmology; thus, the quantity  $\alpha_{\text{iso}}^{\text{rescaling}}$  is telling you how much shift is expected due to **a** blind cosmology. Rescaling the measure values by it, therefore, tells you how far (deflated) you are from **a** blinded cosmology. Moreover, to reduce the sample variance, we focus on the ratio,

$$\Gamma_\alpha = \frac{\alpha^{\text{measured}}}{\alpha^{\text{rescaling}}} \times \frac{1}{\alpha_{\text{baseline}}^{\text{measured}}}. \quad (4.12)$$

The first term accounts for how blinding shifts the measure parameters, while the second one is merely to reduce sample variance.  $\sigma_\Gamma$  are obtained by standard error propagation.

Finally, we generalize the notation defining  $\Gamma_i$ , where  $i$  range in  $(\alpha_{\text{iso}}, \alpha_{\text{AP}}, df)$ , i.e.,  $\Gamma_i = \{\Gamma_{\alpha_{\text{iso}}}, \Gamma_{\alpha_{\text{AP}}}, \Gamma_{df}\}$ . The ratios ( $\Gamma_i$ ) is expected to be unity in an ideal scenario where no systematic errors were introduced. Similarly,  $\tilde{\Gamma}_m$  is defined in terms of differences,

$$\tilde{\Gamma}_m = (m^{\text{measured}} - m^{\text{rescaling}}) - m_{\text{baseline}}^{\text{measured}}. \quad (4.13)$$

The differences ( $\tilde{\Gamma}_m$ ) is expected to be zero in the presence of no systematic errors.

## 5 Validation with mocks

### 5.1 Mock data and preliminary checks

To rigorously validate the blinding scheme developed for DESI DR1, we utilize mock catalogs produced from the ABACUSSUMMIT<sup>14</sup>  $N$ -body simulations [54]. These mocks were produced by fitting the galaxy two-point correlation function at small scales using **Abacus** halos and a halo occupation distribution model [55], in order to populate the dark matter halos with galaxies.<sup>15</sup> Each tracer at each redshift is populated over all 25 base boxes, giving a total volume of  $200h^{-3}\text{Gpc}^3$ . These mock datasets aim to comprehensively mimic the characteristics of the actual DESI DR1 data. Specifically, they include features like target galaxy distributions and redshift bins, observational systematics, and have a window footprint applied, thus providing an ideal dataset for testing our analysis pipeline.

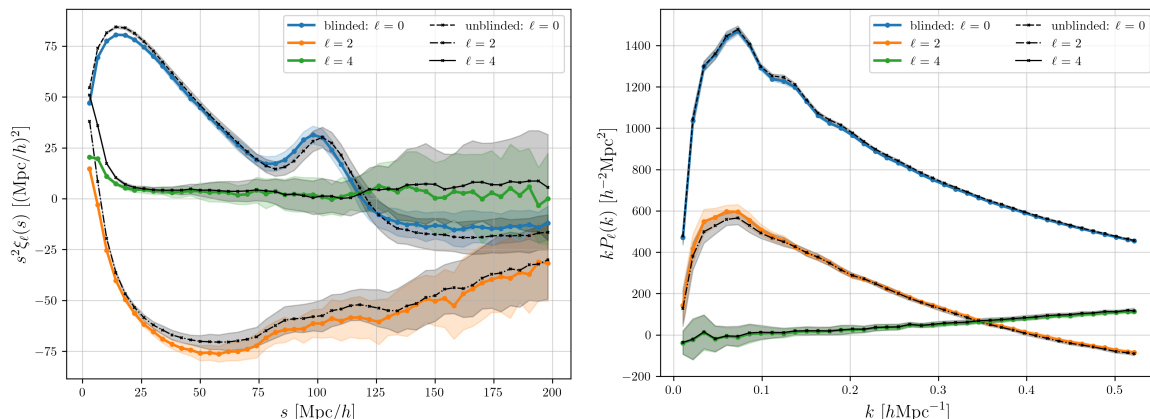
We applied our blinding pipeline to the mock catalogs, using the same procedure as for the actual data, but with a distinct realization of the blinding parameters.

To assess the impact of these blinding parameters, we tested them across 25 ABACUSSUMMIT mock catalogs. We aimed to confirm that the blinding did not inadvertently introduce any distortions or biases in the clustering signal, such as spurious changes to the correlation function or power spectrum. As depicted in figure 2, examining both the correlation function and power spectrum multipoles reveals that the blinded and unblinded monopole amplitudes remain largely consistent, meaning the blinded catalogs retain the same overall shape and comparable amplitudes as the unblinded catalogs, while successfully altering the BAO position and quadrupole amplitude on large scales. This validation serves as a sanity check to confirm that the blinding procedure does not introduce any unwanted artifacts, such as excess clustering on any scale. Therefore, the blinding procedure changes the clustering signal as expected without introducing unexpected distortions or biases.

Second, we select *one* of the 25 ABACUSSUMMIT mock catalogs for more detailed testing, applying  $8 \times 2$  blindings (following the 8 randomly selected  $(w_0, w_a)$  pairs shown as black points in figure 1 and 2  $f_{\text{NL}}$  values,  $\pm 20$ ) to this specific catalog. Note that all our fitting analyses in this section refer to this *one* out of 25 ABACUSSUMMIT mock catalogs we mention earlier. Clustering measurements from this catalog are displayed for completeness in the appendix B.

<sup>14</sup>The ABACUSSUMMIT suite contains 25 different boxes at Planck-2018 cosmology with different phases in the initial conditions, which we refer to as base boxes.

<sup>15</sup>The mocks used in this work are referred to as **Abacus-1** in other DESI papers, i.e., they were produced with the fiducial cosmological parameters of eq. (3.13).



**Figure 2.** Comparison of blinded and unblinded mocks for multipoles  $\ell = 0$ ,  $\ell = 2$ , and  $\ell = 4$ , for the correlation function (left column) and power spectrum (right column). The curves show the mean is across 25 ABACUSSUMMIT catalogs which are blinded with the same blinding parameters.

## 5.2 Testing and evaluation

We carry out several tests on LRGs, ELGs and QSO; see table 1 for relevant details. These tracers are vital for testing the robustness and applicability of our blinding scheme for both BAO and ShapeFit analyses, as these are the samples for which the analysis is planned to be carried out with real data.<sup>16</sup>

Next, we detail the results from the anisotropic pre- and post-reconstruction BAO fitting analysis in both configuration and Fourier spaces, as well as the ShapeFit pre-reconstruction power spectrum analysis.

### 5.2.1 Validation results for BAO analysis

Starting with our first tracer, LRGs, we carry out anisotropic BAO fitting for LRG samples in the redshift ranges of  $0.4 < z < 0.6$ ,  $0.6 < z < 0.8$ , and  $0.8 < z < 1.1$ .

Figure 3 shows the anisotropic BAO fitting results, for both pre- and post-reconstruction. We focus on three metrics: the ratios of measured to expected values for BAO fitting parameters ( $\alpha_{\text{iso}}$ ,  $\alpha_{\text{AP}}$ ), as well as the reduced  $\chi^2$  of the fit;<sup>17</sup> error bars show measurement uncertainties. Here the measured values are using the analysis pipeline, while the expected values are based on the true and blinded cosmology. As we see in the figure, while there is some variation across the 16 blinding catalogs in the measured vs. expected ratios of the BAO fitting parameters,  $\chi^2$  is within  $1\sigma$  for  $\sim 80\%$  of the cases and within  $2\sigma$  for the rest, as represented by the light-gray and dark-gray areas, respectively.<sup>18</sup> Note that these  $\sigma$ -limits are obtained as the standard deviation from the mean of the  $\chi^2$  distribution of the 16 blinding

<sup>16</sup>Note that we are not including BGS and Ly $\alpha$  tracers in our tests; we do, however, have the sample BGS\_BRIGHT-21.5 in the validation tests with real data, discussed in section 6.

<sup>17</sup>We emphasize that the  $\chi^2$  values correspond to the inference of each blinded cosmology, not normalized by the first sim.

<sup>18</sup>There are a few cases near the  $2\sigma$  boundary, but they are not concerning as they remain close. These cases represent only about 3% of the total.

Case	Blinding Parameters			Test Metrics		Fit $\chi^2$
#	$w_0$	$w_a$	$f_{\text{NL}}$	$\Gamma_{\alpha_{\text{iso}}}$	$\Gamma_{\alpha_{\text{AP}}}$	$\chi^2/(52-13)$
1	-0.90	0.03	20	$1.000 \pm 0.0128$	$1.000 \pm 0.0346$	1.56
2	-0.91	-0.68	20	$1.005 \pm 0.0131$	$0.996 \pm 0.0383$	1.26
3	-0.97	-0.51	20	$1.002 \pm 0.0126$	$0.999 \pm 0.0376$	0.90
4	-1.00	0.29	20	$1.003 \pm 0.0128$	$0.986 \pm 0.0358$	2.38
5	-1.05	0.14	20	$1.002 \pm 0.0126$	$1.009 \pm 0.0385$	1.21
6	-1.11	0.45	20	$1.002 \pm 0.0129$	$1.011 \pm 0.0385$	1.22
7	-1.16	0.37	20	$1.005 \pm 0.0135$	$1.003 \pm 0.0408$	0.93
8	-1.23	0.77	20	$1.008 \pm 0.0130$	$1.003 \pm 0.0384$	0.95
9	-0.90	0.03	-20	$1.000 \pm 0.0125$	$0.993 \pm 0.0345$	1.85
10	-0.91	-0.68	-20	$1.003 \pm 0.0128$	$0.996 \pm 0.0375$	1.38
11	-0.97	-0.51	-20	$1.005 \pm 0.0129$	$0.987 \pm 0.0382$	1.17
12	-1.00	0.29	-20	$1.005 \pm 0.0126$	$0.990 \pm 0.0359$	1.48
13	-1.05	0.14	-20	$1.004 \pm 0.0130$	$0.987 \pm 0.0382$	0.89
14	-1.11	0.45	-20	$1.005 \pm 0.0129$	$0.998 \pm 0.0383$	1.04
15	-1.16	0.37	-20	$1.006 \pm 0.0134$	$1.008 \pm 0.0409$	1.08
16	-1.23	0.77	-20	$1.012 \pm 0.0133$	$0.991 \pm 0.0398$	0.96

**Table 3.** Configuration space post-reconstruction anisotropic BAO fitting parameters for the LRG  $0.8 < z < 1.1$  sample from the ABACUSSUMMIT mock catalog, blinded using various blinding cosmologies with varying  $w_0$ ,  $w_a$ , and  $f_{\text{NL}}$ . This table presents two statistics: (1)  $\Gamma_i$ , defined as the ratio of the measured vs. expected values of the  $i$ th parameter (i.e., the two BAO fitting parameters) from each simulation compared to a reference simulation, and (2) the reduced  $\chi^2$  from the fit, comparing the theoretical model to the data. Table 4 presents the numbers for other analysis parameters.

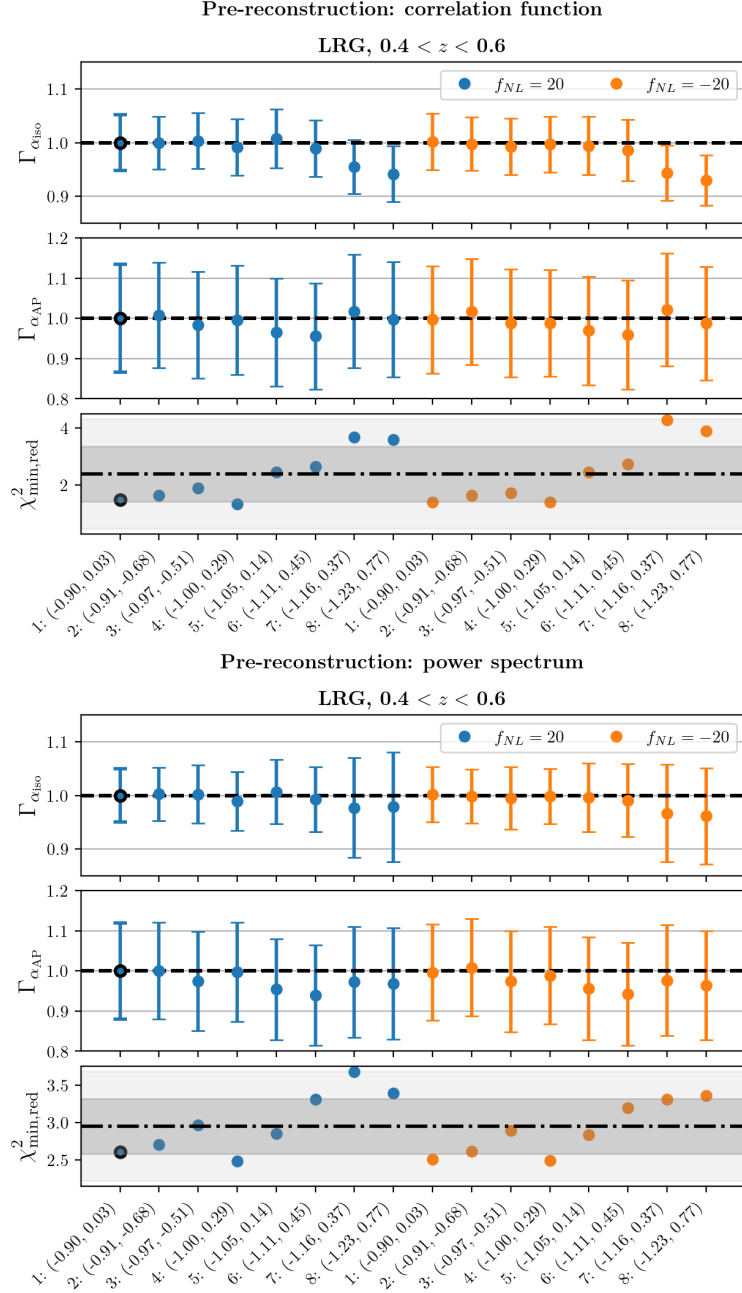
cases. This demonstrates that the blinding preserves the signal we aim to measure, but maintains the variances of the sample  $\chi^2$  low.

Also, it is interesting to note that, as expected, the BAO analysis is not sensitive to the  $f_{\text{NL}}$ -blinding values, i.e., the scatter in  $(\alpha_{\text{iso}}, \alpha_{\text{AP}})$  follows the same trend for  $f_{\text{NL}} = 20$  (blue points) and  $f_{\text{NL}} = -20$  (orange points).

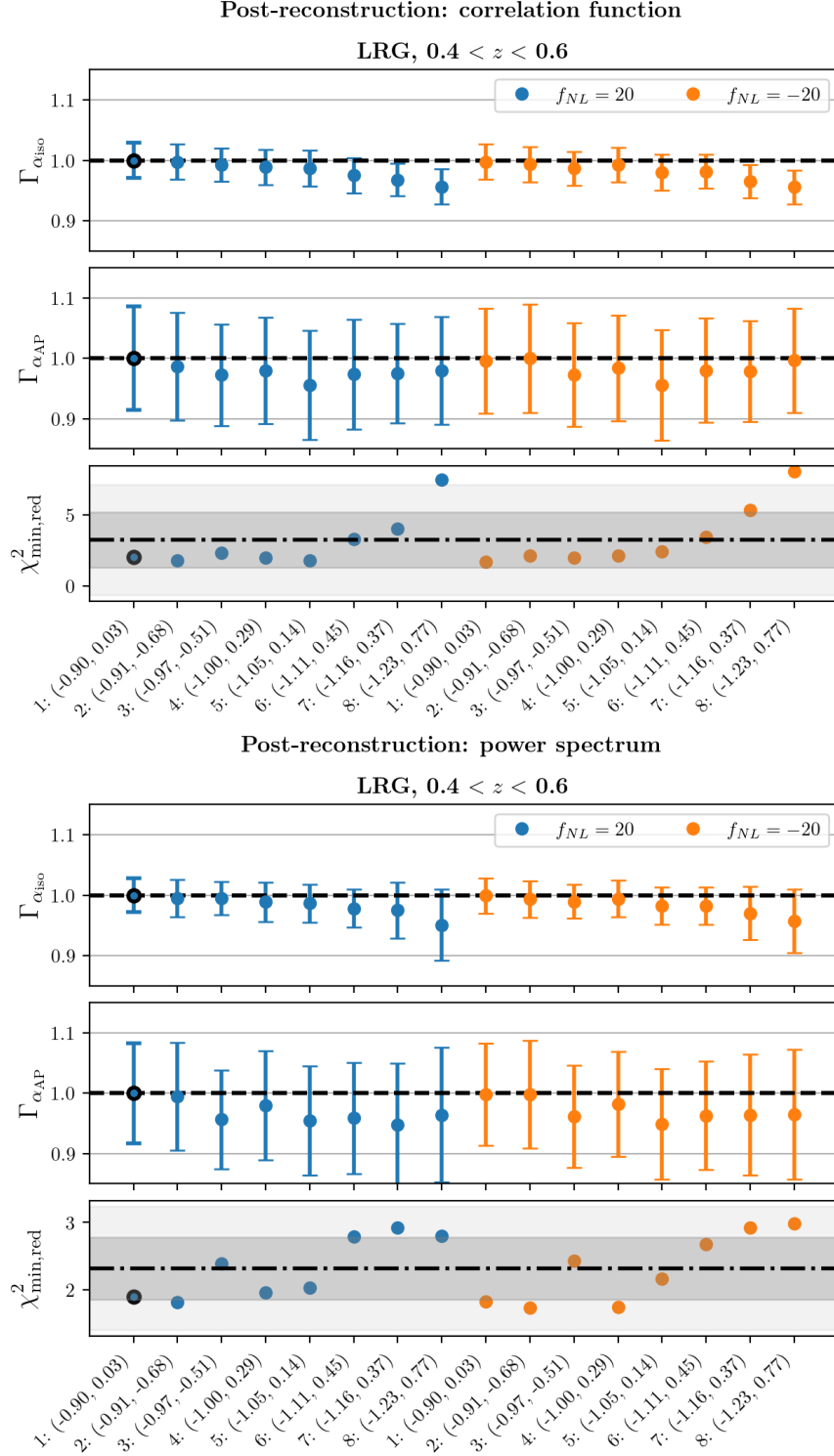
The results of the post-reconstruction anisotropic BAO fitting for the LRG samples, including the quantitative assessments of  $\alpha_{\text{iso}}$ ,  $\alpha_{\text{AP}}$  across the  $8 \times 2$  studied blinded cosmologies, are summarized in table 3. This table provides a comprehensive overview of the fitting accuracy in configuration space, emphasizing that  $\chi^2$  variation is within  $1-2\sigma$ .

We repeat the same for our other two tracers, ELGs and QSO, arriving at the same results: while we see some variation in measured vs expected ratios of the BAO fitting parameters across the 16 blinding catalogs, we see that  $\chi^2$  variation is always within  $1-2\sigma$ ; appendix B shows the figures (figure 14 for pre-construction results for ELGs, figure 15 for post-construction ones; and figure 16 for pre-reconstruction results for QSOs and figure 17 for post-construction ones).

The consistency of our results in both real and configuration spaces underscores the robustness of the blinding technique, which is vital for ensuring unbiased cosmological parameter estimation, as well as the reconstruction pipeline to extract the BAO features. We note that there are a few cases where  $\alpha_{\text{iso}}$  uncertainties is  $\sim 1\sigma$  away from the expected (e.g., pre-recon middle panels of figure 3), but they still consistent within  $2\sigma$ .



**Figure 3.** Pre-reconstruction anisotropic BAO fits using the correlation function (top) and the power spectrum (bottom) for LRG samples for the first redshift bin (each row) from 16 different blinded mock catalogs with  $(w_0, w_a)$  choices identified by indices 1–8 and two  $f_{NL}$  values by blue and orange, respectively. The top two subplots in each panel plot  $\Gamma_i$ , defined as the ratio of measured vs expected ratios of the  $i$ th parameter from each `sim` vs a reference `sim` (identified with black marker-edge); here  $i = \alpha_{iso}, \alpha_{AP}$ , where measured values are from the analysis pipeline while expected ones are from the theoretical connection with the respective  $(w_0, w_a)$ ; error bars capture the measurement uncertainties while propagating the errors. This statistic allows comparing all the sims against a reference sim. The bottom subplot in each panel displays the reduced  $\chi^2$  values, with shaded areas representing  $1\sigma$  and  $2\sigma$  regions; the  $\sigma$ -limits are obtained as the standard deviation from the mean of the  $\chi^2$  distribution of the 16  $\chi^2$  values. This confirms the consistency and reliability of BAO measurements under various blinding shifts given the small variations.



**Figure 4.** Post-reconstruction anisotropic BAO fits for LRG samples for the first redshift bin (each row) following the structure in figure 3. Here, too, we see that while our  $\Gamma$  statistic varies around the expected value of unity, the reduced  $\chi^2$  indicates good fits.

Case	Analysis Hyperparameters											
#	$b$	$d\beta$	$\Sigma_s$	$\Sigma_{\parallel}$	$\Sigma_{\perp}$	$a_{2,0}$	$a_{2,1}$	$b_{0,0}$	$b_{0,2}$	$b_{2,0}$	$b_{2,2}$	
1	$2.103 \pm 0.078$	$1.30 \pm 0.39$	$1.7 \pm 1.8$	$3.4 \pm 1.4$	$2.98 \pm 0.92$	-110	0.13	-0.0004	0.0025	-0.0037	0.0087	
2	$2.05 \pm 0.23$	$1.14 \pm 0.49$	$1.9 \pm 1.9$	$5.1 \pm 1.7$	$2.79 \pm 0.93$	-65	0.13	0.0010	-0.0051	-0.0011	-0.0035	
3	$2.17 \pm 0.22$	$0.87 \pm 0.49$	$1.9 \pm 1.9$	$4.7 \pm 1.7$	$2.70 \pm 0.91$	1.3	0.43	0.00066	-0.0028	-0.00053	-0.0044	
4	$2.093 \pm 0.080$	$1.30 \pm 0.44$	$1.8 \pm 1.8$	$4.5 \pm 1.5$	$2.90 \pm 0.91$	-90.	-0.093	-0.0001	0.0014	-0.0039	0.010	
5	$2.18 \pm 0.22$	$0.80 \pm 0.50$	$1.9 \pm 1.9$	$4.9 \pm 1.7$	$2.70 \pm 0.91$	4.9	0.73	0.0005	-0.0022	-0.0018	-0.0004	
6	$2.05 \pm 0.22$	$1.23 \pm 0.49$	$1.8 \pm 1.9$	$4.9 \pm 1.7$	$2.79 \pm 0.92$	-75	0.062	0.00066	-0.0034	-0.0017	0.0011	
7	$2.01 \pm 0.27$	$1.29 \pm 0.33$	$1.9 \pm 1.9$	$5.8 \pm 1.7$	$2.85 \pm 0.94$	-72	-0.22	0.00059	-0.0027	-0.0004	-0.0014	
8	$2.03 \pm 0.23$	$1.25 \pm 0.56$	$1.9 \pm 1.9$	$5.1 \pm 1.6$	$2.78 \pm 0.93$	-48	-0.29	0.0004	-0.0019	0.0001	-0.0034	
9	$2.049 \pm 0.080$	$1.30 \pm 0.44$	$1.7 \pm 1.8$	$3.6 \pm 1.4$	$2.69 \pm 0.90$	-130	0.21	-0.0011	0.0038	-0.0056	0.018	
10	$2.15 \pm 0.22$	$0.76 \pm 0.57$	$1.9 \pm 1.9$	$4.6 \pm 1.7$	$2.70 \pm 0.92$	37	0.50	-0.0001	-0.00098	-0.0002	-0.0078	
11	$2.186 \pm 0.082$	$0.70 \pm 0.37$	$2.0 \pm 2.0$	$4.9 \pm 1.7$	$2.77 \pm 0.92$	40.	0.68	-0.0002	-0.00078	-0.0005	-0.0050	
12	$2.075 \pm 0.080$	$1.30 \pm 0.48$	$1.7 \pm 1.8$	$4.7 \pm 1.5$	$2.69 \pm 0.91$	-90.	-0.11	-0.0011	0.0049	-0.0038	0.010	
13	$2.08 \pm 0.23$	$1.05 \pm 0.44$	$1.9 \pm 1.9$	$4.9 \pm 1.6$	$2.78 \pm 0.93$	-48	0.43	-0.0004	0.00064	-0.0020	0.0012	
14	$2.10 \pm 0.22$	$0.97 \pm 0.39$	$1.9 \pm 1.9$	$4.9 \pm 1.7$	$2.73 \pm 0.92$	-37	0.56	-0.00057	0.0011	-0.0024	0.0037	
15	$2.03 \pm 0.24$	$1.16 \pm 0.42$	$1.9 \pm 1.9$	$5.6 \pm 1.7$	$2.85 \pm 0.94$	-25	-0.12	-0.0004	0.0001	0.0003	-0.0046	
16	$2.00 \pm 0.24$	$1.26 \pm 0.45$	$1.9 \pm 1.9$	$5.5 \pm 1.6$	$2.79 \pm 0.93$	-77	-0.18	-0.0004	-0.00057	-0.0013	0.0024	

**Table 4.** Values of the various analysis parameters for the blinding cosmologies in table 3 with matching case number #. From left to right: linear galaxy bias ( $b$ ), linear RSD nuisance parameter accounting for the anisotropy of the signal amplitude ( $d\beta = \beta/\beta_{\text{fid}}$  with  $\beta = f/b$ ), Fingers of God damping ( $\Sigma_s$ ), line-of-sight BAO damping ( $\Sigma_{\parallel}$ ), transverse BAO damping ( $\Sigma_{\perp}$ ), with the remaining parameters ( $a_{n,n}-b_{n,n}$ ) being the DESI baseline parametrization for broadband term, capturing any deviation from the linear theory [58].

### 5.2.2 Validation results for ShapeFit analysis

We repeat the framework in section 5.2.1, but now with ShapeFit. We carry out ShapeFit analysis for LRGs in the three redshift ranges of  $0.4 < z < 0.6$ ,  $0.6 < z < 0.8$ , and  $0.8 < z < 1.1$ , with the first redshift shown in figure 5, and other two in the appendix B (figure 18). In addition to  $\alpha_{\text{iso}}$ ,  $\alpha_{\text{AP}}$ , we now have two additional parameters  $m$  and  $df$ , where  $m$  is the additional parameter due to ShapeFit while  $df$  comes from RSD. We see that the measured values of the parameters are close to the expected ones, although with some variation. Again, we note that as in the BAO case  $\alpha_{\text{iso}}$  and  $\alpha_{\text{AP}}$  are insensitive to the  $f_{\text{NL}}$ -blinding choice. However, the ShapeFit parameter  $m$  shows a systematic offset between the  $f_{\text{NL}} = 20$  and  $f_{\text{NL}} = -20$  choices. This is expected, given the degeneracy between the scale-dependent bias and the power spectrum slope on large scales captured by  $f_{\text{NL}}$  and  $m$  respectively [56]. Also, the growth rate  $df$  exhibits a very mild  $f_{\text{NL}}$  dependence, which can be explained by the small correlation of  $df$  with  $m$ . Finally, the  $\chi^2$  of the fits is always within  $1-2\sigma$ , indicating robustness.

We repeat the same for ELGs, arriving at the same results: while we see some variation in measured versus expected parameters across the 16 blinding catalogs, we see that  $\chi^2$  variation is always within  $1-2\sigma$ ; figure 19 shows the results.

By expanding our tests to different tracers and using both BAO and ShapeFit in Fourier and configuration spaces, we have substantially validated the robustness and applicability of our blinding scheme. We refer the reader to [57] for a complete list of unblinding tests and optimizations for the full shape of the power spectrum.



### 5.3 Concluding remarks on validation on mocks

The suite of tests conducted on mock datasets confirms the robustness and efficacy of the blinding scheme developed for DESI DR1. These validation efforts provide strong evidence that our blinding scheme can be reliably used for DR1 and beyond, both for DESI and other large-scale galaxy surveys, mitigating the potential risks of experimenter bias in cosmological parameter inference. It serves as a foundational step toward more complex, multi-probe cosmological analyses that may require intricate blinding techniques.

Specifically, we applied 16 possible blinding configurations to one ABACUS simulation, resulting in 16 new blinded catalogs from which we performed our tests. Based on these tests, we are confident that our blinding scheme does not introduce any unwanted artifacts. While the blinding procedure does alter summary statistics and shift parameter inference, this is a core feature of the blinding process, not a flaw. Importantly, the results demonstrate that the blinding does not degrade the model’s ability to fit the measurement events.

While our tests on mock datasets were comprehensive, it is crucial to note that real-world data might present complexities not accounted for in our mock datasets; we probe these in section 6, where we carry out tests on blinded real data. Future work may include updating the blinding scheme and continually validating it against more complex and realistic mock datasets.

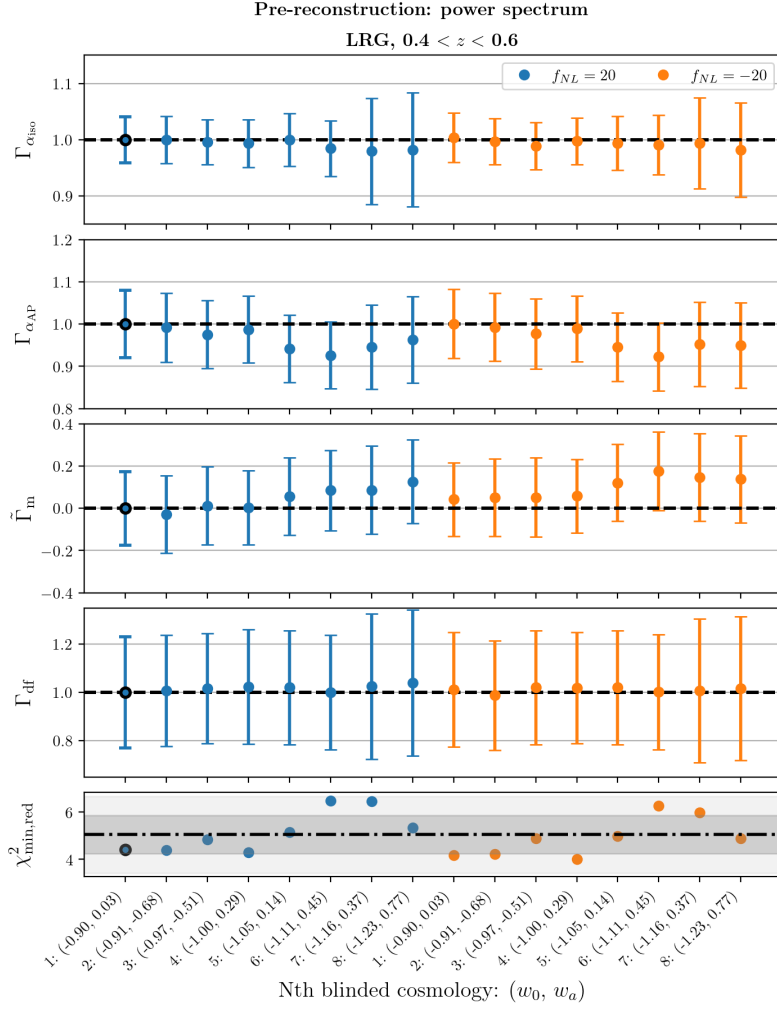
## 6 Validation with real data

So far, we have only discussed validation in the realm of mock datasets. In this section, however, we dive into the validation using real (blinded) data, focusing on only BAO given that constraints from RSD are not unblinded at the time of this writing, as explained in section 3.4. In the following, we detail the methodology we use to further validate the blinding scheme but first, we explain the data we work with.

After the validation tests passed on mocks (as detailed in section 5), we blinded DESI DR1 using the blinding pipeline described in section 3.3, implementing blinding as detailed in section 3.2. The resulting catalog is the *blinded catalog*, consisting of real DR1 data that is blinded. To validate the blinding scheme with real data, we carry out tests on this blinded catalog. As mentioned in section 3.2, we do not yet disclose the parameters used to blind the data.

To probe the efficacy of our blinding scheme, we apply a second layer of blinding to the blinded catalog, using a  $(w_0, w_a)$  pair that is at the edge of the allowed region in order to stress-test the scheme while arbitrarily choosing  $f_{\text{NL}} = 20$ . This second layer of blinding enables us to directly check that our blinding strategy does not introduce any unintended biases and that our analysis pipelines are robust, as we can apply and then remove this second layer at will.

For clarity, we refer to the original blinded catalog as the *fiducial blinded catalog* while the doubly-blinded data is referred to as the *double-blinded catalog*. Comparing analysis results from the fiducial blinded and double-blinded catalogs allows us to check the impact of blinding while preserving characteristics of real data that may not have been captured fully in the mocks.



**Figure 5.** ShapeFit fits using LRG samples for the first redshift bin (each column) from 16 different blinded mock catalogs. Various details here are the same as in figure 3, except that  $i = \alpha_{\text{iso}}, \alpha_{\text{AP}}, df, m$  in  $\Gamma_i$  while  $\tilde{\Gamma}_i$  is the same as  $\Gamma_i$  but comparing differences as opposed to ratios between measured and expected (since expected is 0). As for BAO fits, we see that the ratios (differences) are close to 1 (0) and the  $\chi^2$  variations are within  $1-2\sigma$ , demonstrating the robustness of the fits.

## 6.1 Validation tests and results

### 6.1.1 Varying analysis choices

As a first test, we run the baseline analysis pipeline, defined below, on the double-blinded catalogs, alongside a suite of alternative fitting choices:

- **Baseline:** we follow the baseline configuration adopted for the anisotropic BAO analysis ( $\alpha_{\text{iso}}, \alpha_{\text{AP}}$ ) defined in [58]. This consist of a configuration space analysis of post-reconstruction catalogs, using the **Rec-Sym** convention for reconstruction [59], spline-based broadband parameterization, and Gaussian priors on redshift-space distortion parameters ( $\Sigma_s, \Sigma_{\parallel}, \Sigma_{\perp}$ ).

- 1D fit: as a test, we run the baseline analysis except that we perform an isotropic BAO analysis ( $\alpha_{\text{iso}}$ ).
- Pre-recon: as a test, we run the baseline analysis on pre-reconstruction data as opposed to the post-reconstruction data.
- Power spectrum: as a test, we run the post-reconstruction analysis but in Fourier space.
- Polynomial broadband: as a test, instead of using the spline parameterization for the broadband, we employ a polynomial.
- Flat priors: as a test, instead of using informative priors on redshift-space distortion parameters, we use flat priors.

Overall, our comparative analysis across these configurations yields consistency with the baseline results, notwithstanding minor variations discussed below. Figure 6 demonstrates these effects, showcasing the parameter constraints for each galaxy tracer under different analysis settings.

The 1D fit closely mirrors baseline findings, with deviations within  $1\sigma$  for most tracers. We see that reconstruction markedly enhances precision across most tracers, illustrating the utility of the technique in sharpening parameter estimates. The outlier of this trend is the tracer QSO which tends to be shot-noise limited and therefore does not benefit strongly from sharpening techniques [60].

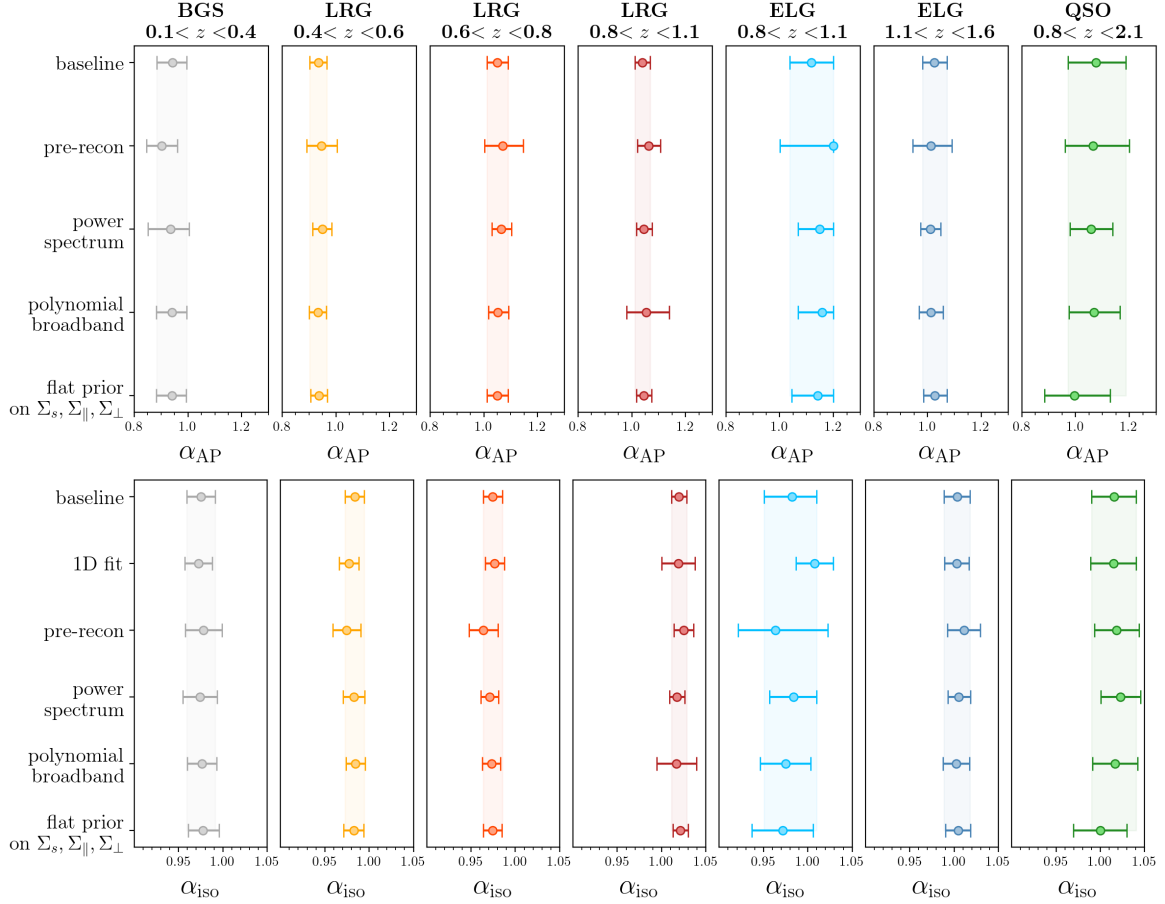
Analysis in Fourier space yields results compatible with those from configuration space, with deviations remaining within  $1\sigma$  for most tracers. This consistency extends to the comparison between spline-based and polynomial broadband parameterizations; the former, a new spline-basis parameterization proposed by [58] while the latter used in BOSS [61] and eBOSS [62].

Lastly, employing flat priors for BAO damping parameters is largely consistent with the baseline, using Gaussian priors, affirming the robustness of our analytical approach to prior effects.

We have demonstrated that the double-blind tests are robust against variations in the fitting procedures. Similarly, the *single-blind* analysis has been thoroughly examined in the companion DESI paper (see, for example, figure 15 in [59]), to which we refer without duplicating the results here to avoid redundancy. Together, these two tests — single-blind and double-blind — confirm that the choice of fitting procedure is not influenced by the blinding technique.

### 6.1.2 Comparing posteriors

Once we establish that our baseline analysis is robust against changes in various choices, we proceed with another test: a comparison of parameter estimates from the fiducial blinded catalog vs. the double-blinded one. Given that we know the blinding parameter used for blinding the double-blinded catalog, we can generate *shifted fiducial* estimates, whereby we shift the inferred values from the fiducial catalog. This mimics a posterior-level blinding, achieved by simply multiplying the fiducial posterior ( $X$ ) by the expected shift ( $a_{\text{shift}}$ ) of

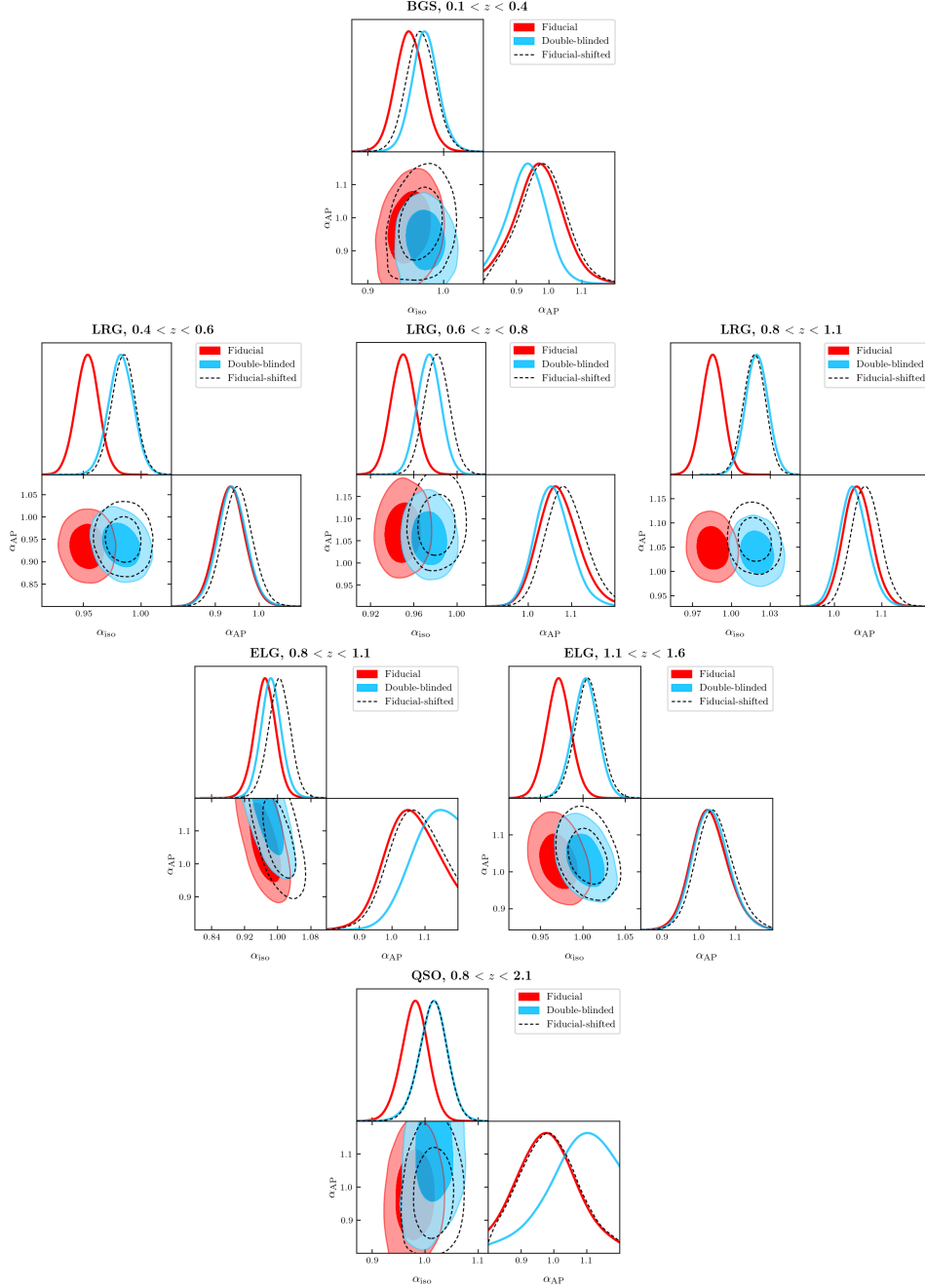


**Figure 6.** Comparison of BAO fitting parameter ( $\alpha_{\text{AP}}$ ,  $\alpha_{\text{iso}}$ ) measurements from the double-blinded catalog, using the baseline analysis model and various fitting choices across the four galaxy tracers (BGS, LRG, ELG, and QSO tracers) for different redshift ranges. The top panel shows the whisker plots for  $\alpha_{\text{AP}}$ , while the bottom panel shows those for  $\alpha_{\text{iso}}$ . We see that the various fitting choices are consistent with the baseline.

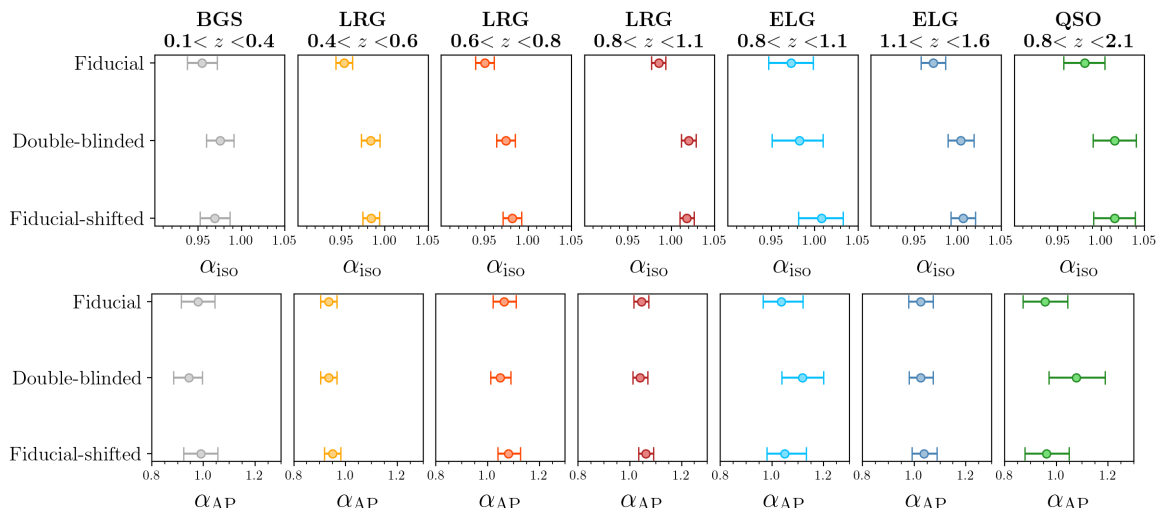
the parameters due to blinding, leading to a fiducial shifted posterior,  $Y_{\text{shifted}} = a_{\text{shift}} X$ ; we calculate  $a_{\text{shift}}$  using the cosmology used for (second) blinding and its relation with the BAO fitting parameters (as presented in eq. (3.2)).

Figure 7 shows the posteriors for the four tracers, across various redshift bins. We see that the fiducial-shifted curves (black lines) do not perfectly match with those from double-blinded (blue), highlighting the distinction between a catalog-level blinding vs. a posterior-level one, i.e., the catalog-level blinding acts on the recovered parameters broadly as expected, but is more complex and thus not identical to simply shifting the posteriors. This figure also demonstrates that the catalog-level blinding works in that the inferred parameters from the double-blinded catalog (blue) do not match those from the fiducial one (red).

As a summary, figure 8 shows the whisker plots for the best-fit values for the two BAO parameters. We see that the fiducial-shifted estimates do not always follow those from the double-blinded catalogs, reinforcing the distinct impacts of a catalog-level blinding as opposed to a posterior-level one; here, the difference between the fiducial cosmology used



**Figure 7.** Post-reconstruction anisotropic BAO fits different tracer samples for fiducial blinded and double-blinded catalogs. Each subplot presents the 68% and 95% confidence level marginalized posteriors for the isotropic ( $\alpha_{\text{iso}}$ ) and anisotropic ( $\alpha_{\text{AP}}$ ) scaling parameters. The red contours denote the posteriors based on the fiducial blinded catalog, while the blue contours represent those from the double-blinded catalog. The dashed contours are produced by applying the same offset to the fiducial posteriors that was used in the blinding process — a check to understand how the data-blinding prescription affects the posteriors. For the cases where the dashed-black posteriors overlap with blue ones, the blinding scheme essentially has the same effect as it would if we have blinded at the posterior level. However, we see that in most cases, the two posteriors are not exactly the same, meaning that our blinding is more complicated than a posterior-level blinding (as is indeed the case). It is also reassuring that the double-blinded posteriors do not match the fiducial ones, indicating that our blinding scheme is effective at blinding for the parameters of interest.



**Figure 8.** Comparison of BAO fitting parameter ( $\alpha_{\text{iso}}$ ,  $\alpha_{\text{AP}}$ ) measurements from the fiducial blinded and double-blinded catalogs, alongside fiducial-shifted which mimics posterior-level blinding. As in figure 6, we consider all four tracers and plot the two parameters in the two rows. We see that fiducial-shifted estimates do not always follow those from the double-blinded catalogs, reinforcing the distinct impacts of a catalog-level blinding as opposed to a posterior-level one, while demonstrating that our blinding scheme effectively masks the parameters of interest.

for the (double) blinding vs. the true (blinded) cosmology underlying the fiducial blinded catalog plays a role. Nevertheless, we find that our blinding scheme is effective at blinding the underlying cosmology, as it should.

## 6.2 Concluding remarks on validation on real data

The tests presented above provide a robust framework for evaluating the impact of blinding on our ability to extract cosmological information. Notably, the consistent results across different fitting methodologies affirm the resilience of our blinding scheme against analytical variations. This exercise also reinforces our confidence in the blinding process and the reliability of subsequent cosmological interpretations. As a result, we set the stage for applying similar methodologies to future data releases from DESI and other large-scale structure surveys.

## 7 Conclusions

In modern observational cosmology, it is crucial to employ blinding methods to safeguard the results against experimenter bias. In this work, we presented and validated a comprehensive blinding scheme for the DESI DR1 analysis, aimed at mitigating experimenter bias and ensuring the integrity of our cosmological parameter estimation. Through a series of rigorous tests on both mock datasets and actual blinded data, we demonstrated the effectiveness of our blinding strategy in preserving the statistical properties of the data while preventing inadvertent unblinding or bias introduction by the researchers.

Our validation process spanned a variety of tracers and included two compression techniques, i.e. BAO and ShapeFit. In particular, we performed a BAO analysis on all

dark-time DESI tracers (see section 5.2.1 for the first redshift bin LRG, i.e.,  $[0.4, 0.6]$  and appendix B for the LRG  $([0.6, 0.8], [0.8, 1.1])$ , ELG and QSO samples) for 16 different blinding configurations of the **abacus-1** mock. Furthermore, we performed the BAO analysis both in configuration and Fourier space, either pre- or post-reconstruction. In all cases, we found exquisite agreement (better than  $1 - \sigma$ ) of the BAO scaling parameters with the expectation. We applied the ShapeFit methodology on the same set of 16 blinded LRG (and ELG) mocks in section 5.2.2 (and appendix B) and found all ShapeFit parameters to agree with the expectation within  $1 - \sigma$ . Finally, in section 6, we performed a series of tests of the BAO pipeline on blinded and double-blinded data for all DESI tracers, finding our baseline choice to be robust against choosing variations such as pre-recon, power spectrum, and 1D fits or adopting a different number of broadband terms or different priors on the BAO damping parameters; we were able to largely recover the fiducial blinded posterior, by shifting the double-blinded posterior. To our knowledge, this is the first time a blinding scheme was explicitly tested with such a doubled layer.

The development and successful validation of this blinding scheme marked a significant step in ensuring that our analysis of DESI DR1 is free from experimenter bias. Furthermore, the methodologies and insights gained from this work offer valuable lessons for future DESI data releases as well as other large-scale structure surveys.

We note that there are analyses that are beyond the scope of current work, including especially the full-shape modeling of the power spectrum; [12] will address this, including blinding validation with full-shape modeling. As we look ahead, the validated blinding scheme will serve as a crucial component of our analysis toolkit, enhancing the credibility of our findings and strengthening the foundation of cosmological research. Future studies will benefit from this foundational work, providing a stepping stone for the application of rigorous scientific methodologies in the exploration of our Universe.

## 8 Data availability

The data used in this analysis will be made public as part of DESI Data Release 1. Details can be found in <https://data.desi.lbl.gov/doc/releases/>.

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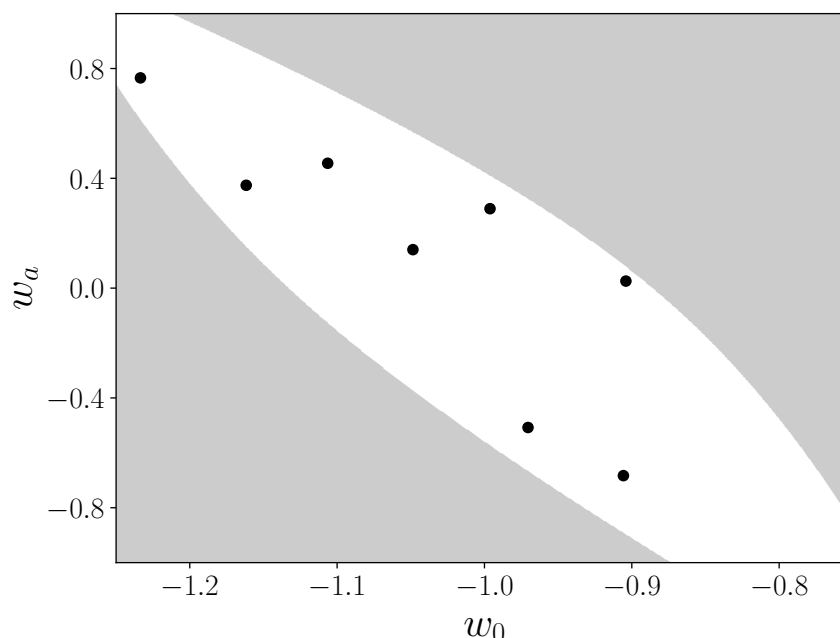


Science and Technology Facilities Council of the United Kingdom; the Gordon and Betty Moore Foundation; the Heising-Simons Foundation; the French Alternative Energies and Atomic Energy Commission (CEA); the National Council of Humanities, Science and Technology of Mexico (CONAHCYT); the Ministry of Science and Innovation of Spain (MICINN), and by the DESI Member Institutions: <https://www.desi.lbl.gov/collaborating-institutions>. Any opinions, findings, and conclusions or recommendations expressed in this material are those of the author(s) and do not necessarily reflect the views of the U.S. National Science Foundation, the U.S. Department of Energy, or any of the listed funding agencies.

The authors are honored to be permitted to conduct scientific research on Iolkam Du'ag (Kitt Peak), a mountain with particular significance to the Tohono O'odham Nation.

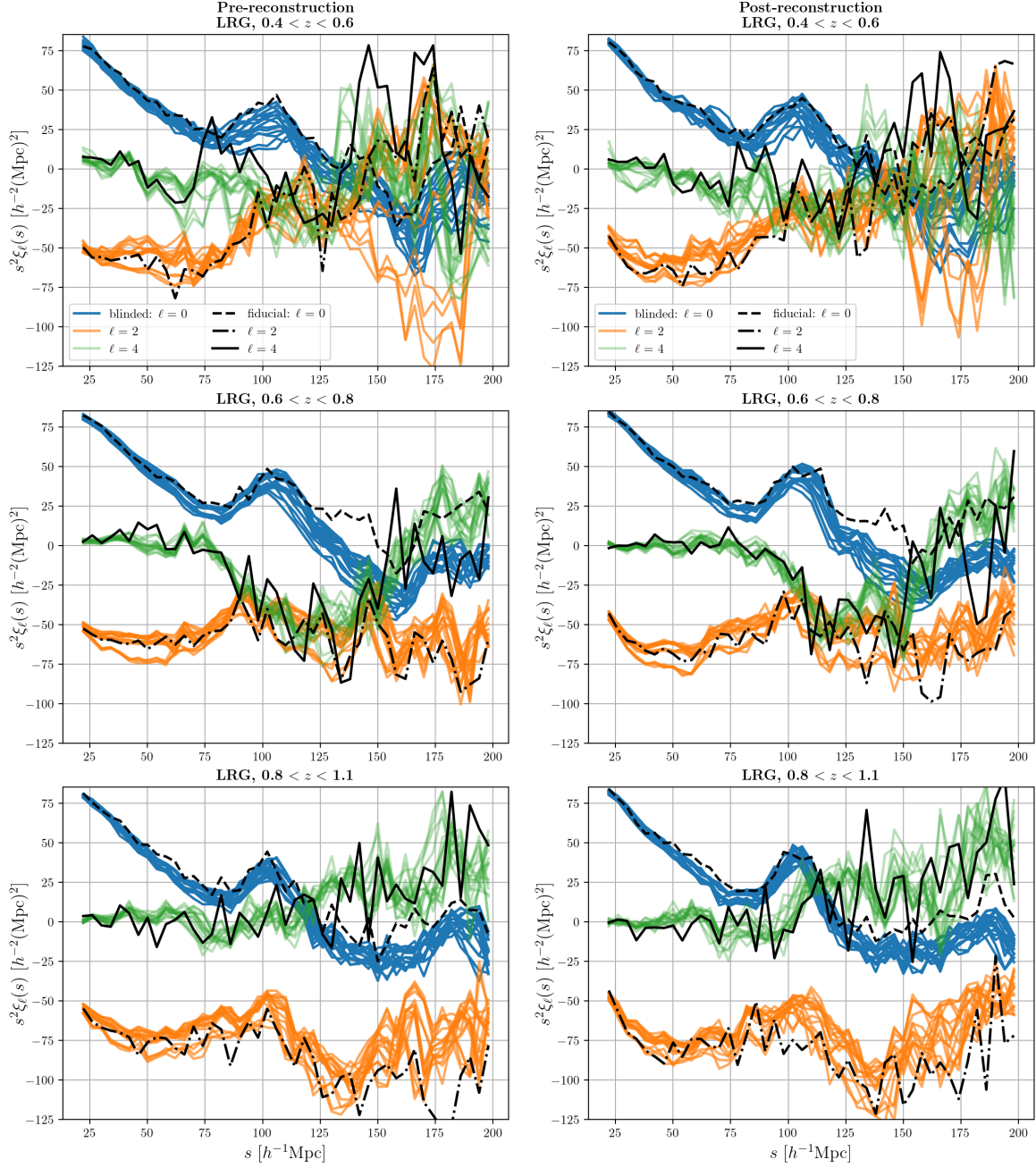
## A Redshift range for blinding parameter space

At the very early stages of the analysis, the redshift range chosen to define the region allowed for  $w_0$ ,  $w_a$  (the white region in figure 1) was thought to encompass all dark-time tracers. Given that LRGs span redshifts 0.4 to 1.1, ELGs 0.8 to 1.6, and QSOs 0.8 to 2.1, we chose  $0.4 < z < 2.1$  as our default redshift range. At a later stage, we also included the BGS in the blinding pipeline, which spans redshift 0.1 to 0.4. However, we decided not to modify the redshift range used to create figure 1 because we found that including the redshift range  $0.1 < z < 0.4$  had very little impact. In figure 9 we show the same plot as in figure 1 but using  $0.1 < z < 2.1$  instead of  $0.4 < z < 2.1$ . We find that the two figures are identical.

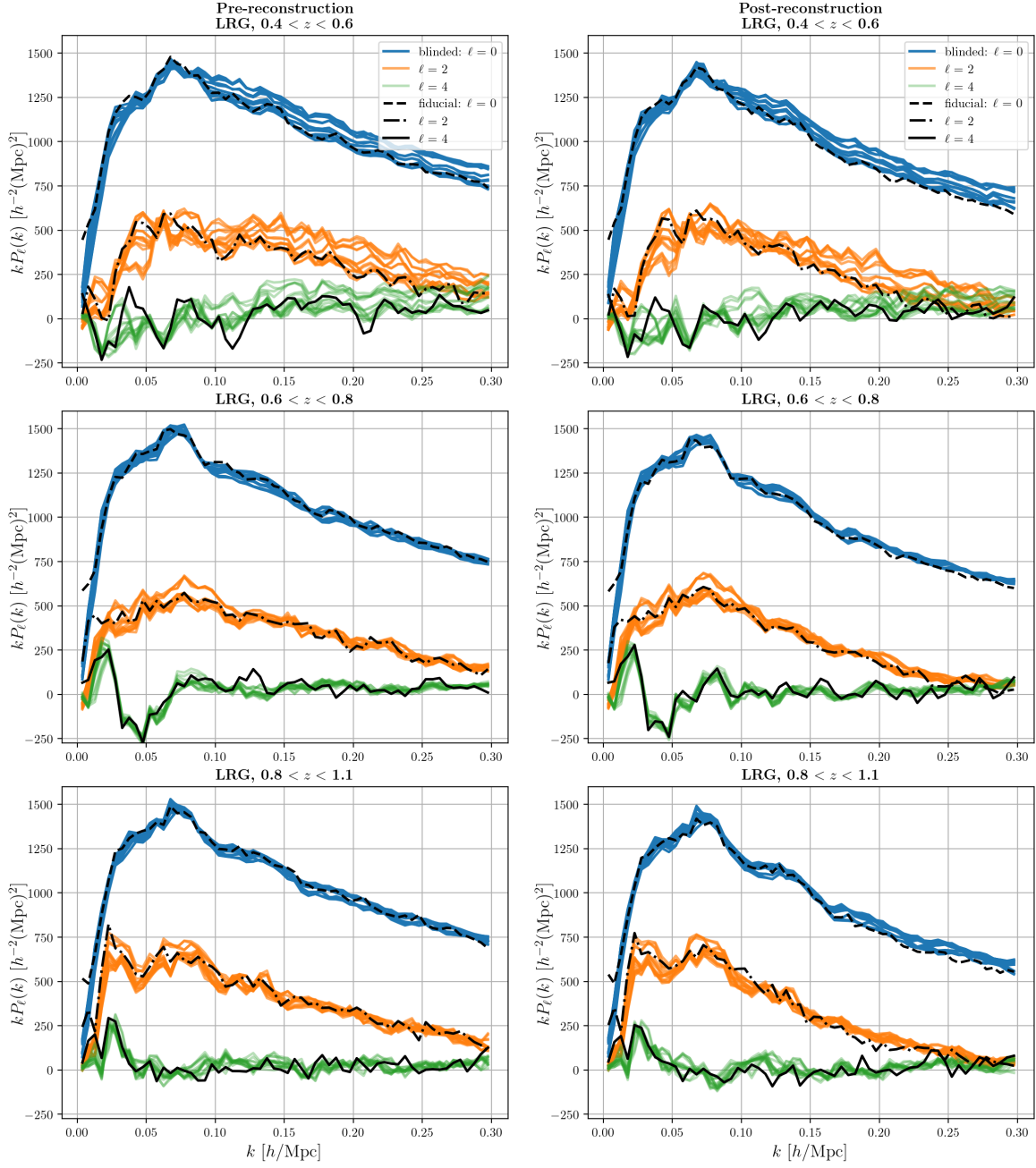


**Figure 9.** Same as figure 1 but using the redshift range  $0.1 < z < 2.1$  to define the allowed parameter space. We see no differences.

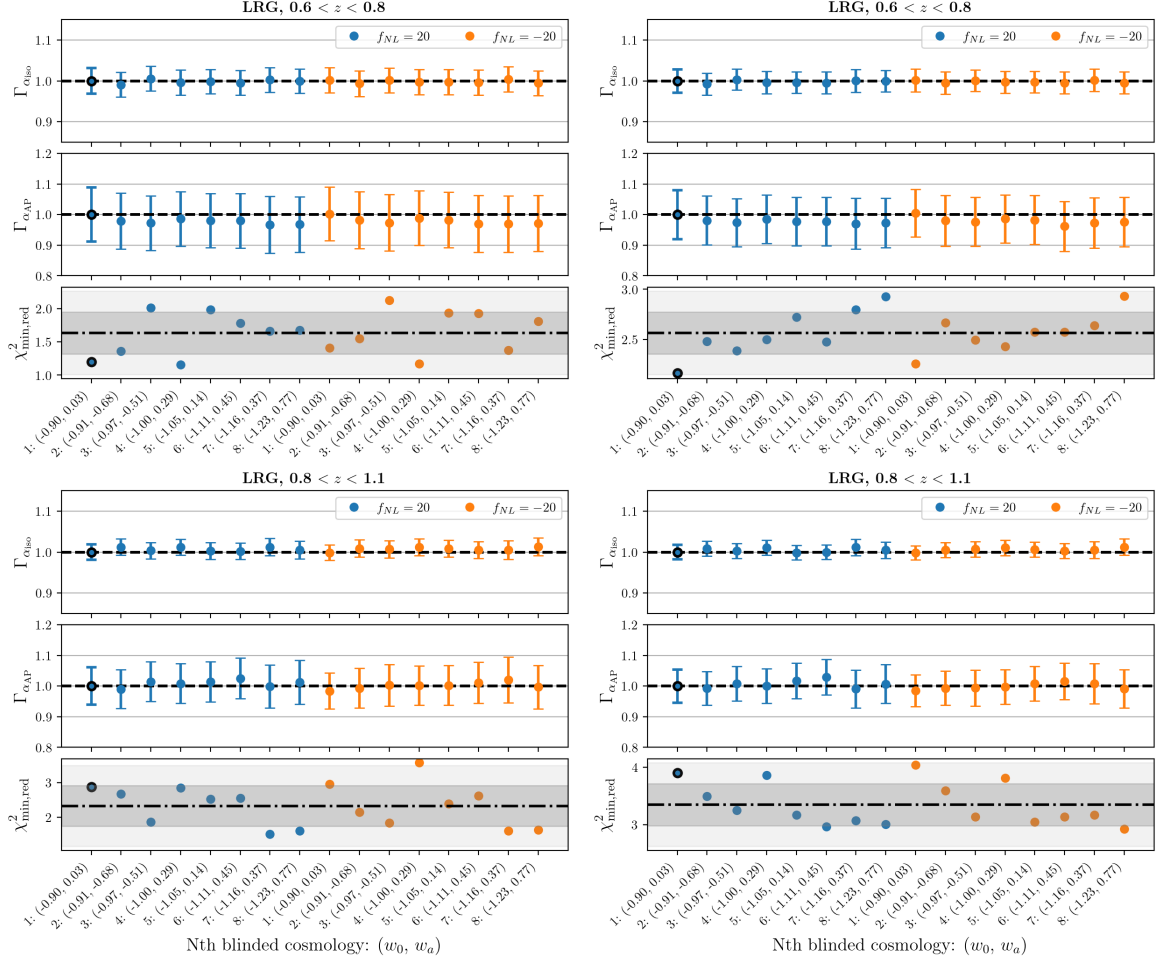
## B Supplemental plots



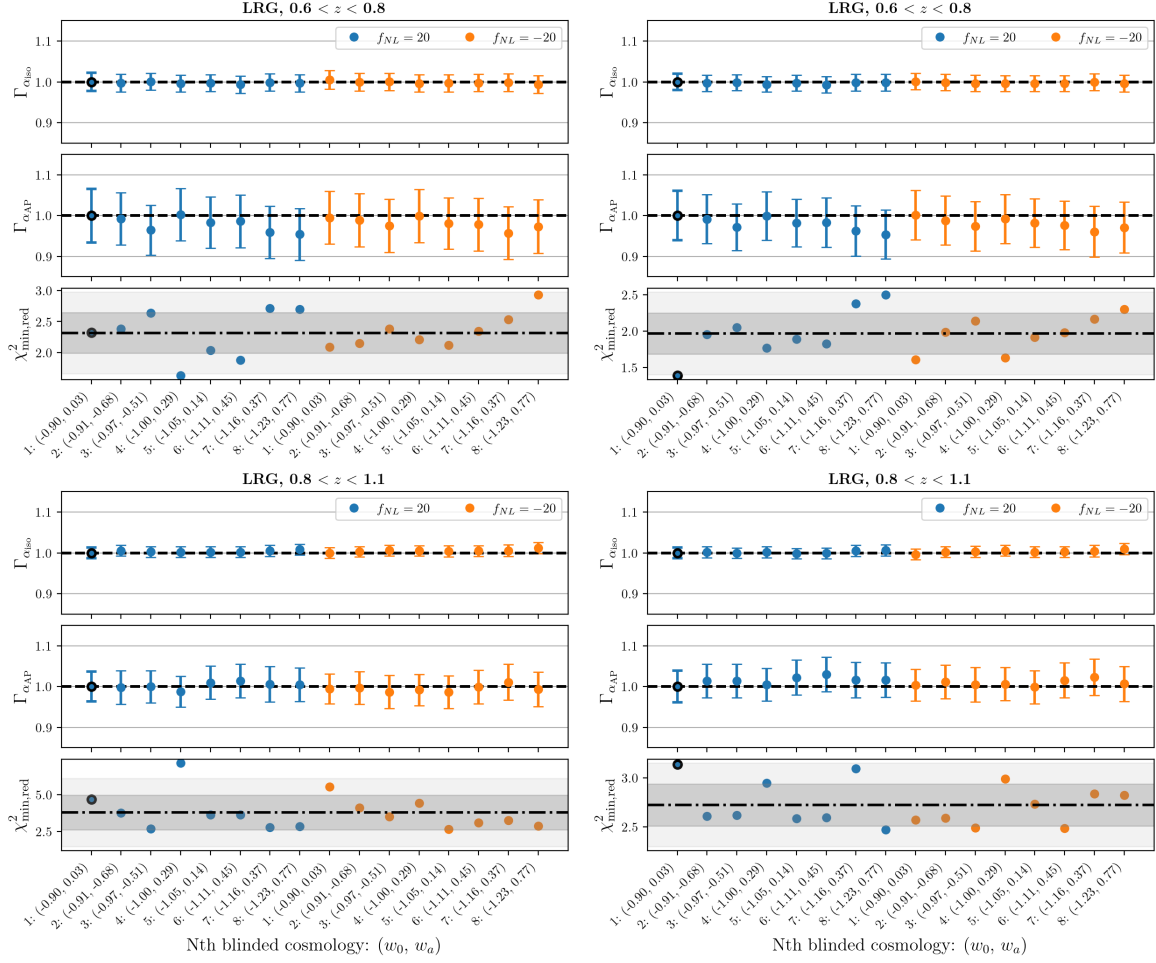
**Figure 10.** Correlation function multipoles for LRG samples from an ABACUSSUMMIT mock catalog, presented for three redshift bins; pre-reconstruction multipoles (left) and post-reconstruction ones (right). The black lines show the fiducial simulation which serves as the baseline for generating 16 different blinded cosmological configurations; these are depicted by the colored lines, with the three colors showing the three multipoles (blue for monopole ( $\ell = 0$ ); orange for quadrupole ( $\ell = 2$ ); green for hexadecapole ( $\ell = 4$ )); all panels use the same legend. These configurations are used throughout the paper in order to assess the impacts of different blinding cosmological configurations on our inference.



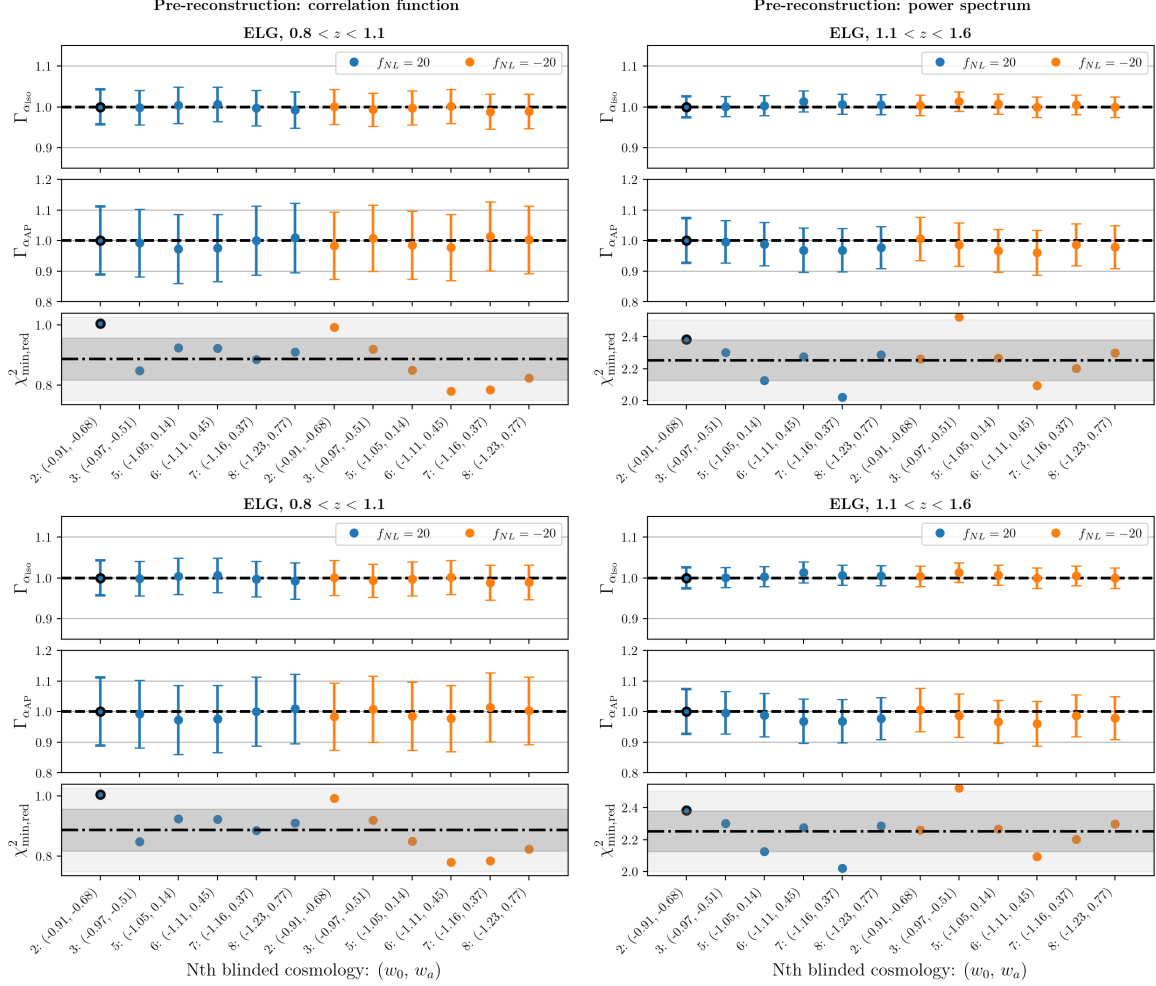
**Figure 11.** Power spectrum multipoles for the same LRG sample as in figure 10. The rows are for different redshift bins; pre-reconstruction multipoles (left) and post-reconstruction ones (right); see figure 10 caption for the rest of the details.



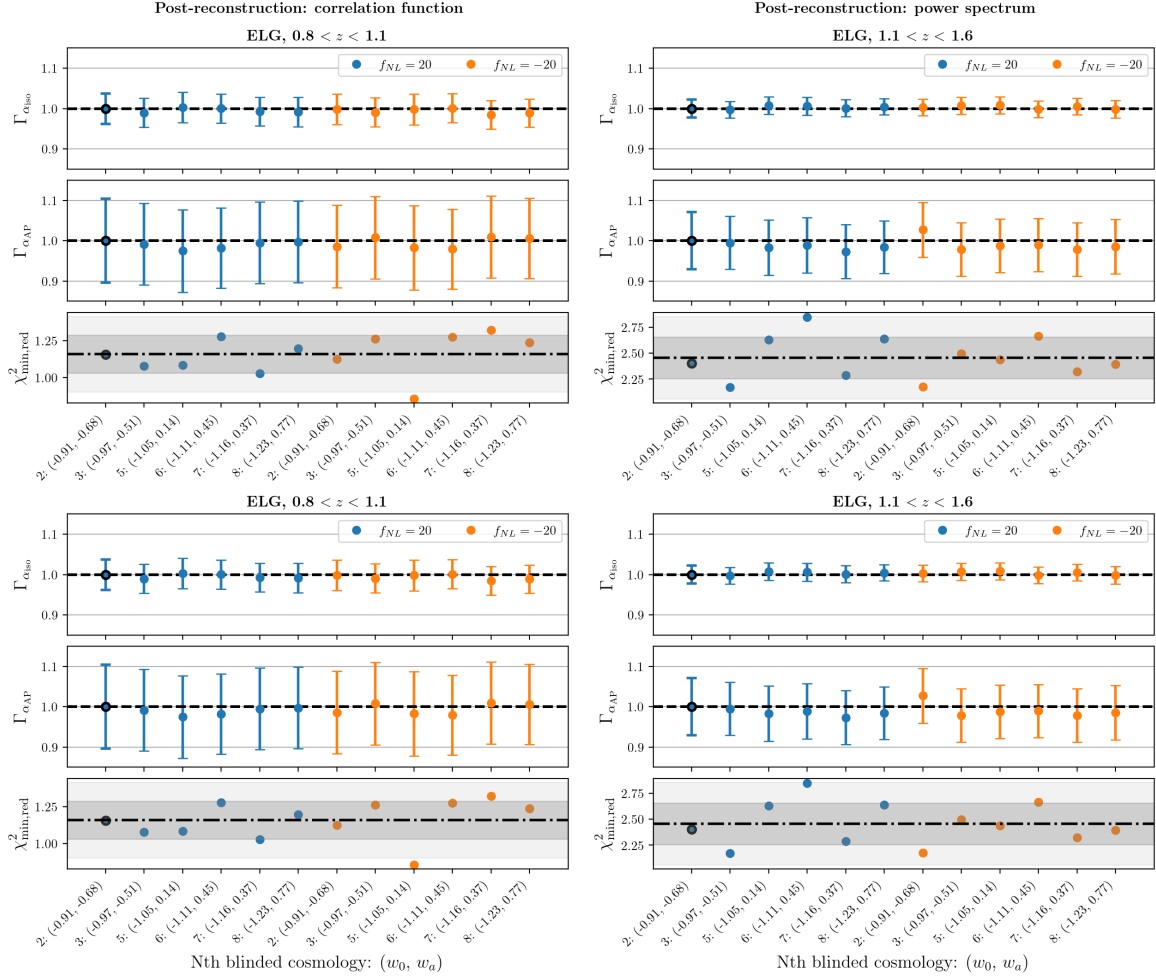
**Figure 12.** Pre-reconstruction anisotropic BAO fits using the correlation function (left column) and the power spectrum (right column) for LRG samples for the two redshift bins (each row) from 16 different blinded mock catalogs with  $(w_0, w_a)$  choices identified by indices 1–8 and two  $f_{NL}$  values by blue and orange, respectively. The top two subplots in each panel plot  $\Gamma_i$ , defined as the ratio of measured vs expected ratios of the  $i$ th parameter from each `sim` vs a reference `sim` (identified with black marker-edge); here  $i = \alpha_{\text{iso}}, \alpha_{\text{AP}}$ , where measured values are from the analysis pipeline while expected ones are from the theoretical connection with the respective  $(w_0, w_a)$ ; error bars capture the measurement uncertainties while propagating the errors. This statistic allows comparing all the sims against a reference sim. The bottom subplot in each panel displays the reduced  $\chi^2$  values, with shaded areas representing  $1\sigma$  and  $2\sigma$  regions; the  $\sigma$ -limits are obtained as the standard deviation from the mean of the  $\chi^2$  distribution of the 16  $\chi^2$  values. This confirms the consistency and reliability of BAO measurements under various blinding shifts given the small variations.



**Figure 13.** Post-reconstruction anisotropic BAO fits for LRG samples for the two redshift bins (each row) following the structure in figure 3. Here, too, we see that while our  $\Gamma$  statistic varies around the expected value of unity, the reduced  $\chi^2$  indicates good fits.

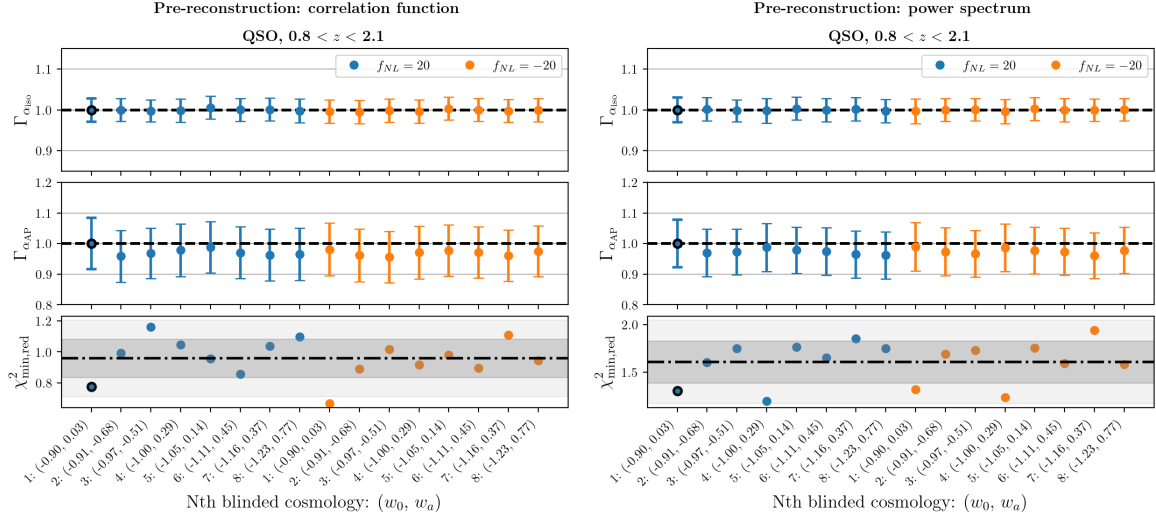


**Figure 14.** Pre-reconstruction anisotropic BAO fits for ELG samples for two different redshifts, following the structure in figure 3. Here too, we see that the measured vs. expected ratios of the BAO fitting parameters vary a little across the sims and the reduced  $\chi^2$  indicates good fits. Note that for this tracer, we drop two of the blinded cosmologies ( $(w_0, w_a)$  pairs 1,4) since these ELG catalogs had few randoms and therefore did not deliver reliable clustering measurements.

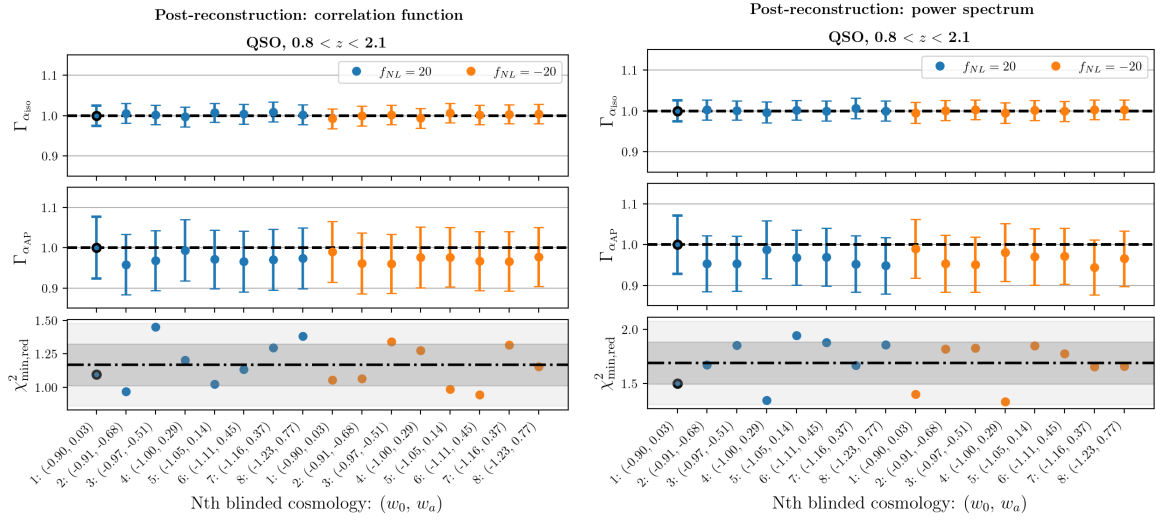


**Figure 15.** Post-reconstruction anisotropic BAO fits for ELG samples, following the structure in figure 3. Here too, we see that the measured vs. expected ratios of the BAO fitting parameters vary a little across the sims and the reduced  $\chi^2$  indicates good fits. As mentioned in figure 14, we drop two of the blinded cosmologies  $((w_0, w_a)$  pairs 1,4).

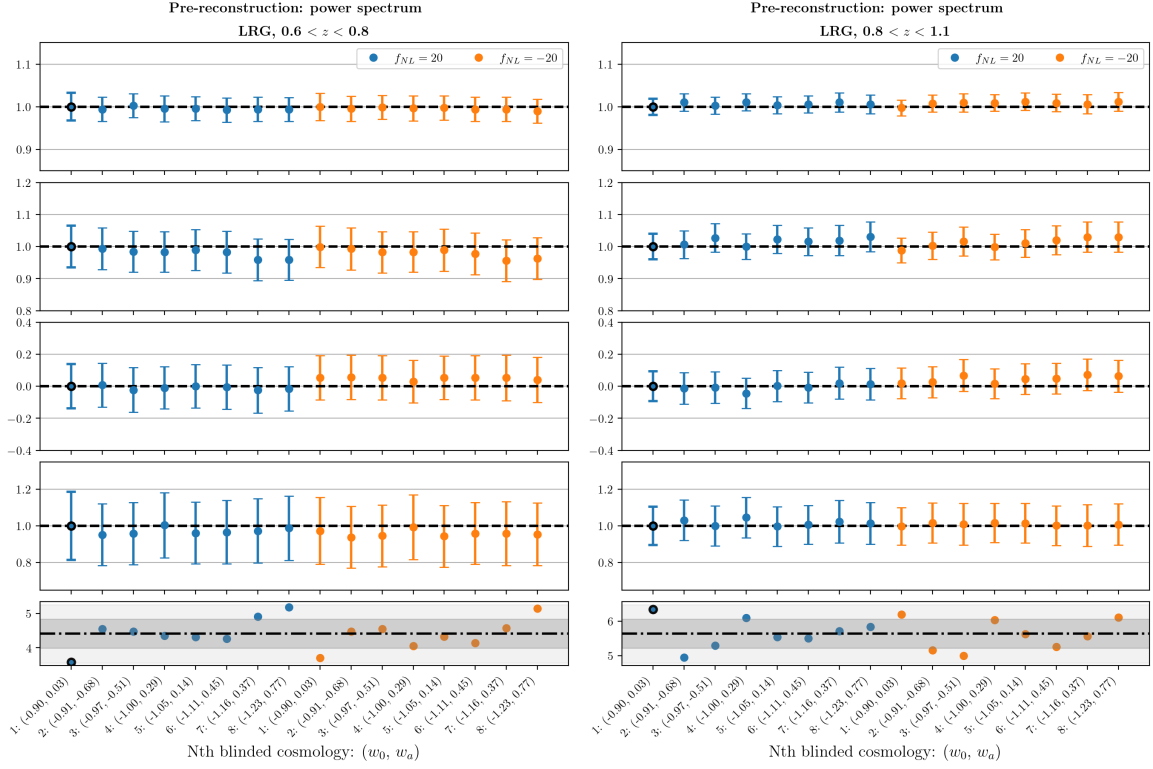




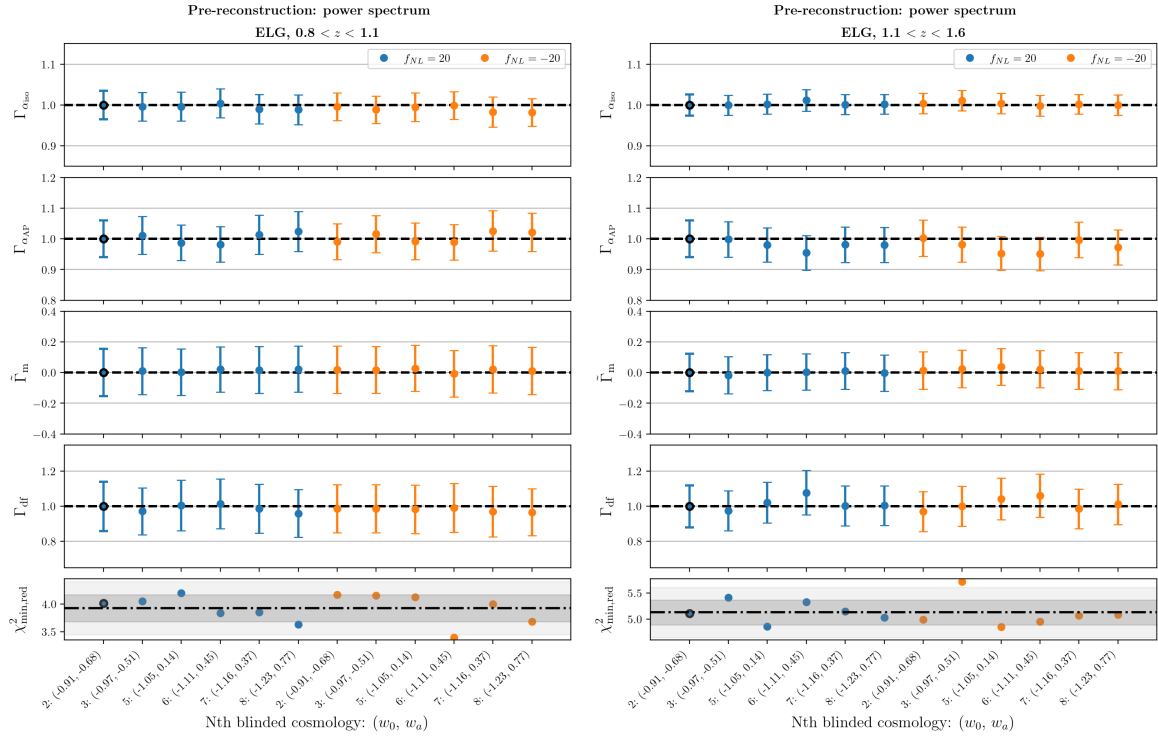
**Figure 16.** Pre-reconstruction anisotropic BAO fits for a QSO sample, following the structure in figure 3. Here too, we see that the measured vs. expected ratios of the BAO fitting parameters vary a little across the sims while the reduced  $\chi^2$  indicates good fits.



**Figure 17.** Post-reconstruction anisotropic BAO fits for a QSO sample, following the structure in figure 3. Here too, we see that the measured vs. expected ratios of the BAO fitting parameters vary a little across the sims and the reduced  $\chi^2$  indicates good fits.



**Figure 18.** ShapeFit fits using LRG samples for the two redshift bins (each column) from 16 different blinded mock catalogs. Various details here are the same as in figure 3, except that  $i = \alpha_{\text{iso}}, \alpha_{\text{AP}}, df, m$  in  $\Gamma_i$  while  $\tilde{\Gamma}_i$  is the same as  $\Gamma_i$  but comparing differences as opposed to ratios between measured and expected (since expected is 0). As for BAO fits, we see that the ratios (differences) are close to 1 (0) and the  $\chi^2$  variations are within  $1\text{--}2\sigma$ , demonstrating the robustness of the fits.



**Figure 19.** ShapeFit fits for ELG samples, following the structure in figure 18. Here, too, we see that while the measured vs. expected ratios/differences of the various parameters vary a little across the sims, the reduced  $\chi^2$  variations are within  $1\text{--}2\sigma$ . As mentioned in figure 14, two sims are dropped for this tracer.

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